

DISCUSSION PAPER SERIES

IZA DP No. 14536

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## ABSTRACT

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# The Impact of Sleep Restriction on Interpersonal Conflict Resolution and the Narcotic Effect\*

Insufficient sleep is commonplace, and understanding how this affects interpersonal conflict holds implications for personal and workplace settings. We experimentally manipulated participant sleep state for a full week prior to administering a stylized bargaining task that models payoff uncertainty at impasse with a final-offer arbitration (FOA) procedure. FOA use in previous trials decreases the likelihood of voluntary settlements going forward—the narcotic effect. We also report a novel result that a significantly stronger narcotic effect is estimated for more sleepy bargaining pairs. One implication is that insufficient sleep predicts increased dependency on alternatives to voluntarily resolution of interpersonal conflict.

**JEL Classification:** J52, D74, D90, C92, D83

**Keywords:** bargaining, sleep restriction, arbitration, dispute/conflict resolution, narcotic effect

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## INTRODUCTION

There has been an increased focus recently on sleep and workplace outcomes such as health, productivity, and conflict (Barnes and Watson, 2019; Robins et al, 2019; Simon et al, 2020). Recent data from the National Health Interview Survey indicates a significant increase in working U.S. adults with insufficient sleep levels from 2010 to 2018 (Khubchandani and Price, 2020). It therefore seems wise to investigate the variety of ways in which sleepy workers affect workplace outcomes, which also holds implications for effective managerial strategies. The objective of this paper is to examine how insufficient sleep impacts interpersonal conflict in a stylized zero-sum bargaining environment. Workplace conflict alone is estimated to cost millions of working days annually in time lost resolving frictions between individuals within an organization.<sup>1</sup> Conflict management is a primary function of managers that requires time and energy (Mintzberg, 1973; Brahnam et al, 2005), and efforts to successfully deal with workplace conflict often distinguish more versus less successful managers (Luthans et al, 1985). Understanding the interaction between conflict resolution and sleepiness is also important outside of the workplace (e.g., commercial disputes and divorce settlement), and therefore our findings have practical implications for both personal and occupational settings.

Adverse health outcomes or productivity loss are known to result from insufficient sleep. In the U.S. alone, recent survey evidence indicates that sleep loss may cost employers over a million working days each year (Hafner et al, 2017), which translates to tens of billions of dollars in lost productivity.<sup>2</sup> The importance of having well-rested employees in an organization extends beyond health and direct productivity effects (see Litwiller et al, 2017; Barnes and Watson, 2019). Conflict is costly because it detracts from other responsibilities, it lowers morale and motivation, and results in an overall less productive and enjoyable home or working

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<sup>1</sup> CPP, Inc, publisher of the Myers-Brigg Type Indicator® (MBTI) assessment, produced the CPP Global Human Capital report (2008) based on results from survey data commissioned in partnership with business psychology firms in foreign countries to survey 5000 full-time employees in nine countries from Europe and the Americas (see [https://img.en25.com/Web/ CPP/Conflict\\_report.pdf](https://img.en25.com/Web/ CPP/Conflict_report.pdf) accessed October 22, 2020). Their data showed that, on average, 2.1 hours per week were spent dealing with conflict of some sort by each employee. For the U.S. this translates to well over 300 million working days per year, but their findings indicated the proportion of working hours lost in dealing with conflict each week was even higher in other countries (e.g., Germany and Ireland). While the type of conflict we consider in this study is only a subset of what the reports include, it is likely that insufficient sleep contributes to less observable opportunity costs of conflict that remain unmeasured and absent in the literature.

<sup>2</sup> Proportionally similar losses due to insufficient sleep are reported for Germany, the UK, Japan, and Canada (Hafner et al, 2017), and such estimates are focused on lost productivity and health-related costs.

environment. For these reasons there are significant opportunities to improve both interpersonal relationships and the company's bottom line with the right strategy that understands the relationship between sleep and conflict resolution.

Regarding workplace environments where conflict may arise, it has been shown that insufficient sleep lowers an employee's organizational commitment and job satisfaction (Barnes, 2011). Insufficient sleep also negatively affects interpersonal relationships, work-team effectiveness, and long-term work relationships (Barnes and Watson, 2019), which may result at least partly from a reduced propensity of workers to resolve interpersonal conflict themselves. It is clear that sleep is a workplace health concern. Workplace health promotion programs (WHPPs) are not uncommon in U.S. organizations, but their focus is typically on weight loss and smoking cessation (Robbins and Girardin, 2018). Our research suggests a wider scope of benefits from sleep-focused WHPPs than previously considered, and it adds to a growing body of literature focused on how sleep impacts interpersonal behaviors.<sup>3</sup>

We study a zero-sum bargaining environment that uses a specific procedure to model how uncertainty is resolved when settlement efforts fail: final-offer arbitration (FOA). In our experiment, pairs of individuals bargain over the value of an asset, which is analogous to any outcome of value where interests may compete. If they cannot voluntarily reach an agreement, a third-party (the arbitrator) resolves the conflict. The arbitrator draws randomly from a distribution of values to help settle disputes either in favor of one individual or the other. While formal FOA benefits have been previously noted in the literature (Wall and Callister, 1995), managers often perform *informal* FOA to resolve workplace conflicts (Nugent and Broedling, 2002). Even more generally, the most important feature of FOA for our purposes is that the procedure introduces uncertainty regarding one's payoff or outcome in the event of unresolved conflict. Given this feature, its application extends to many environments. For example, allowing any conflict to go unresolved also implies uncertainty regarding the ultimate payoff one will enjoy (i.e. one's household or workplace utility, in a general sense). While arbitration in our bargaining environment has no direct monetary cost (thus, no explicit resource losses), unresolved disputes are implicitly costly because individuals generally prefer to craft their own

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<sup>3</sup> Other relevant findings should be noted as well. Wagner et al (2012) linked sleep loss to increased "cyberloafing" in the work place, Barnes et al (2011) documented increased unethical behaviors due to poor sleep (as rated by workplace supervisors), and Barnes et al (2015) found that poor sleep quality led to abusive supervisory behaviors and poor work unit engagement.

settlement (Crawford 1979). Additionally, workplace conflict may explicitly impact the productivity of those in conflict or even pull other workers (or managers) away from productive activities to help manage the conflict. While these costs are not explicit in our experiment, our results imply that sleepiness can cause these transactions costs to increase as sleepy individuals rely more heavily on third-parties to resolve conflict. The reduced propensity to voluntarily resolve conflict by sleepy individuals is, we suggest, related to how sleep may promote overconfidence (or optimism) regarding the uncertain conflict outcome, as well as how the addictive nature of arbitration-type options to resolve conflict may be moderated by sleepiness.

The methodology we used to examine sleep was an ecologically valid sleep protocol with randomized treatment assignment, a full week of sleep-restricted (SR) or well-rested (WR) sleep levels, and passive but objective sleep data measurements. Sleep levels were experimentally manipulated to either: (i) a level commonly experienced but deemed insufficient and a public health problem (CDC, <https://www.cdc.gov/sleep/index.html>) for the young adults we study or (ii) to a level considered sufficient and recommended by health experts for young adults (National Sleep Foundation, [www.sleepfoundation.org](http://www.sleepfoundation.org)). In this way, our participants are in a manipulated sleep state when administered the zero-sum bargaining environment with a Final-Offer Arbitration (FOA) procedure when there is failure to voluntarily settle the conflict. The present study contributes novel evidence to examine how common levels of insufficient sleep impact settlement likelihood when conflict outcomes are uncertain, and our use of an anonymous and structured laboratory bargaining environment can help identify causation with less difficulty than would be the case with naturally occurring field data.

## **BACKGROUND**

Chronic sleep restriction (i.e.,  $\leq 6$  hr/night) is commonplace for 30% or more of U.S. adults (Schoenborn and Adams, 2010), and such levels of insufficient sleep have been documented in several industrialized countries and estimated to cost these economies anywhere from 1% to 3% of their annual GDPs (Hafner et al, 2017). According to these estimates, the workplace is responsible for much of these losses due to estimated productivity losses (i.e., increased absenteeism or *presenteeism* among those sleeping less than 6 hours per night). Certain occupational constraints may contribute to this. Irregular work hours such as shift work, which is performed annually by over 20 million U.S. wage and salary workers (McMenamin, 2007),

and commute times are linked to reduced nightly sleep. The impact of insufficient sleep is certainly not limited to the workplace or reduced productivity concerns. Increased conflict, which is a key focus in our paper, is costly to interpersonal relationships, workplace settings, and society as a whole. Yet, little is known about how insufficient sleep impacts outcomes in social environments or how it affects conflict propensity in bargaining.

There exists some recent research that has started to examine the impact of sleep loss on socially interactive environments, which we consider relevant to our study of bargaining and conflict. The general finding points to reduced pro-sociality when sleep deprived or restricted (see Anderson and Dickinson, 2010; Dickinson and McElroy, 2017; Holbein et al, 2019). In other research that relates sleep to interpersonal behaviors, poor sleep quality impacts power dynamics between supervisors and subordinate work units (Barnes et al, 2015), is detrimental to conflict resolution among couples (Gordon and Chen, 2013), and it seems to decrease one's empathy (Guadagni et al, 2014). Kahn-Greene et al. (2006) explored emotional intelligence and responses to depictions of frustrating situations meant to examine how one may direct blame or aggression towards others. Their study involved a relatively small sample (n=26) of active duty military following extreme total sleep deprivation (TSD) of 55 hrs, but their findings suggest that sleep loss likely impacts bargaining outcomes in ways that are often difficult to quantify. Specifically, they observed that sleep loss increased one's tendency to blame others and to exhibit hostility in frustrating situations. A recent review of the literature also recommends employees be trained in effective sleep hygiene after concluding that sleep issues in the workplace can contribute to aggressive responses and workplace deviance, both of which are likely important antecedents to interpersonal conflict (Budnick and Barber, 2015). As a whole, these previous studies start to paint a picture of how sleepy individuals may impact interactions and conflict in both workplace and personal environments. Our study of bargaining behavior and conflict propensity among well-rested versus sleepy individuals contributes to this body of literature.

A main interest of ours involves bargaining in the shadow of uncertainty regarding impasse outcomes, which we mimic with the use of a binding arbitration procedure that has been shown to be susceptible to optimism or overconfidence (Dickinson, 2006). An important question, therefore, is whether sleep deprivation promotes increased optimism in such settings. While we know of no direct evidence that supports the proposition that insufficient sleep leads to

overconfidence or optimism in bargaining, Venkatraman et al. (2011) identified changes in economic preferences following TSD that correlated with neural activation pattern changes of interest for this question. Specifically, emotion center activation following monetary losses was moderated and activation in prefrontal regions responsive to monetary gains increased following TSD, which manifested behaviorally as increased risk taking.<sup>4</sup> Others have also reported that more mild *partial* sleep restriction or suboptimal circadian timing increase monetary risk taking (Maric et al, 2017; Castillo et al, 2017), and extreme adverse sleep states should not be the only concern regarding relevant behavioral outcomes. Though a bit more speculative, we interpret the findings that TSD enhances deliberative-thought activation regarding gains, and reduces emotional-thought activation when experiencing monetary loss, as setting the stage for optimism in bargaining under conditions of more commonly experienced sleep restriction.

Sleep loss may produce a decreased sensitivity to losses or unfavorable outcomes in conflict, even in more mild adverse sleep environments. While we are not arguing that chronic partial sleep restriction effects on decision making will mimic those found under TSD, some research supports this possibility to at least some degree. For example, certain neurobehavioral impairments found with TSD have been documented under conditions of chronic partial sleep as well (Van Dongen et al, 2003).<sup>5</sup> Also, recent research reporting reduced prosocial behavior (Anderson and Dickinson, 2010) or Bayesian decision tendencies (Dickinson and Drummond, 2008) due to TSD were also found when using an at-home chronic partial sleep-restriction protocol more similar to the present study (Dickinson and McElroy, 2017, and Dickinson et al, 2016, respectively). We interpret this to mean that insufficient sleep at more moderate levels may also promote optimism in bargainers who become relatively more focused on the prospect of gains or self-serving settlements compared to the possibility of loss.

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<sup>4</sup> A larger literature has found increased risk taking with TSD, in general, although one of the more carefully designed studies to examine pure risk taking articulated its conclusion by stating that TSD decreases *sensitivity* to monetary risk (McKenna et al, 2007).

<sup>5</sup> The manipulation of sleepiness bears some resemblance also to research on cognitive load manipulations. While beyond the scope of this paper, traditional cognitive load manipulations that involve digits recall or other memorization prior to the task of interest need not produce similar behavioral results to the negative impact on cognition known to result from sleep loss. Deck and Jahedi (2015) present new evidence but also survey much of the literature on cognitive load and economic decision tasks. Among other less relevant findings, they report that cognitive load is typically found to increase risk aversion but have mixed effects on generosity that may be relevant to bargaining and conflict resolution. These findings do not necessarily match those of studies administering sleep deprivation or restriction, which generally find that sleepiness increases risk taking behaviors, and decreases generosity and pro-social behaviors as noted earlier in the paper.

Importantly, sleep deprivation has also been linked to general decrements in the type of prefrontal brain activation necessary for handling the unexpected, revision of strategy, or effective communication (see review in Harrison and Horne, 2000). Thus, sleepy bargainers with an available conflict resolution procedure may find it simpler to rely on such procedures rather than putting in the effort to negotiate a voluntary settlement. In the workplace, for example, sleepy bargainers may then come to develop reliance on third-party assistance as a type of new “status quo” way to resolve conflict (or, they simply become more willing to accept the uncertain outcome from unresolved conflict). In this case, sleepiness in the workplace is not just a productivity or health concern, but it also becomes relevant as a somewhat hidden contributing factor to the successful resolution of interpersonal conflict.

While our research examines the link between sleep and optimism in bargaining, the extant literature is quite clear that optimism or overconfidence in bargaining is common and there is a systematic relationship between overconfidence and dispute or bargaining outcomes.<sup>6</sup> Several early studies documented the importance of bargainer overconfidence in predicting conflict (i.e., failure to negotiate a settlement) using controlled mock negotiations (Bazerman and Neale, 1982, 1983; Neal and Bazerman, 1985). Babcock and co-authors (Babcock and Loewenstein, 1997; Babcock et al, 1997) further documented this and included evidence from field data involving school board vs teachers’ union negotiations. Such research clearly reveals that bargainers who are overconfident regarding their bargaining position or the likely outcome from unresolved conflict are more likely to experience failed negotiations. Uncertainty with respect to payoffs or utility in the event of unresolved conflict may help breed optimism. In the case of formal arbitration, the procedure may guarantee a settlement but possibly decrease *voluntary* settlements (Dickinson, 2006) or even create dependence on the conflict resolution procedure (see Dickinson, 2020, for more on this point). More generally, individuals may grow accustomed to accepting the risky outcome that arises when conflict goes unresolved rather than making efforts to resolve the conflict themselves.

While one’s inclination towards preferring unresolved conflict outcomes can grow stronger in a general sense, previous research on the “narcotic” effect in a bargaining context has

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<sup>6</sup> Others have documented *under*-confidence in the domain of individual performance (see Clark and Friesen, 2009), but our interest is not in this area of beliefs regarding one’s performance. Rather, our interest is in the classic area of self-serving bias whereby individuals’ viewpoints tend to skew in the direction that favors their own interests.

focused on bargainers becoming addicted to a specific procedure. The extant literature is less clear on the existence of a narcotic effect of formal arbitration on dispute resolution. Wheeler (1975) suggested the presence of a narcotic effect but was also cautious about the possibility of other unobserved effects being an issue in field data. Butler and Ehrenberg (1981) concluded, in fact, that once one controlled for unobserved heterogeneity there was actually an estimated *negative* narcotic effect, where past use of arbitration would decrease the likelihood of using arbitration in the present.<sup>7</sup> Champlin et al. (1997) also reported a negative narcotic effect, using 1970s field data on public sector labor disputes, and others have concluded that any positive narcotic effect would be short lived and disappears after a few years (see Chelius and Extejt, 1985; Kochan, 2010). On the other hand, Currie (1989) reported a strong empirical finding from Canadian public sector data that using arbitration in the past made a bargaining unit at least 10% more likely to use arbitration in the present round of negotiations. As such, the field data evidence is mixed regarding the existence of a narcotic effect of arbitration, and these data consider bargaining over group outcomes (i.e., union workers versus management) where collective efforts of bargaining agents are put into each offer and counter-offer on behalf of the principals.

Our research speaks more to (micro-level) interpersonal conflict that one may experience on a regular basis at the home or workplace, and so results from previous laboratory research on bargaining and the narcotic effect are particularly relevant. Controlled laboratory methods can provide a useful complement to field data in such instances, because the bargaining environment used to generate data is free from many of the confounds that complicate field data analysis. Bolton and Katok (1998) concluded from their laboratory study that arbitration slowed the rate of learning about the other party's preferences, which they considered a type of narcotic effect. A more traditional dependency effect of commonly used arbitration mechanisms was reported in the experimental study of Dickinson (2005), where results showed that a more frequent past history of using arbitration in a zero-sum bargaining environment significantly decreased a pair's likelihood of voluntarily resolving conflict in the current bargaining round.<sup>8</sup> More generally,

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<sup>7</sup> When we use the term "narcotic effect" by itself, we mean to say a *positive* narcotic effect, where past use of arbitration increases the likelihood of using arbitration in the present.

<sup>8</sup> Others have used laboratory methods to study bargaining under similar arbitration rules but in somewhat different environments than ours. Deck and Farmer (2007) examined an environment where one bargainer received a fixed payment while the other is the residual claimant of the uncertain value over which they bargain. Another example is Pecorino and Van Boening (2001), who implemented a laboratory environment that allowed

such results suggest that past history of unresolved conflict, when this implies an *uncertain* outcome, may lead to an increased preference for the risky outcome that stems from failed settlement efforts.

## **METHODS AND MATERIALS**

Our experimental design used a two-player zero-sum bargaining environment introduced by Ashenfelter et al. (1992) and with additional adaptations following Dickinson (2004; 2006). In the present study, our key manipulation was to vary the level of sleep restriction of the bargainers. In total, 233 participants completed the sleep protocol and participated in the bargaining experiment.<sup>9</sup> Our starting point for recruiting participants was a large online database we maintain, which we refer to as the “sleep survey database”. This database was maintained and updated regularly by sending new invitations to lists of randomly selected student emails from the investigators’ institution. The recruiting database is unique in the sense that it contains information on self-reported sleep patterns and disorders, mood disorders, diurnal preference as well as more traditional demographic information. Subjects in the database complete an initial prescreening online sleep survey, and from their responses we were able to determine the viability of participants for recruitment to the main study. We would only recruit those who were not at significant risk of major depressive or anxiety disorders, who did not self-report a sleep disorder or insomnia, who did not have extreme diurnal preferences (i.e., extreme morning-type or evening-type preferences based on a validated instrument within the prescreening survey), and who were within the 18-40 year old young adult age range.<sup>10</sup>

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the transmission of private information after the submission of final offers that would be binding under the arbitration rules. Dickinson (2020) cites additional references of laboratory arbitration studies with still different objectives, such as comparisons across different arbitration rules.

<sup>9</sup> A few pairs were completed by use of participants who we recruited for the decision experiment session, but who did not take part in the sleep protocol (n=9 participants). These subjects are classified neither as sleep restricted (SR) nor well-rested (WR) in our analysis. Because of the high fixed cost of each participant who completed the sleep protocol, and our need to have even numbers of participants in each cohort to run the decision experiments, we utilized some of these backup participants to complete pairs. The backup participant decisions are not analyzed in the individual-level analysis. When a bargaining pair contains a backup participant, the pair is excluded from the analysis when conditioning on compliance. We include such pairs in the full-sample “intent-to-treat” analysis and consider such participants as neither SR=1 nor WR=1 when scoring the indicator for the number of SR participant in the pair. That said, intent-to-treat specifications that include pair-level measures of Epworth sleepiness are dropped from analyses because we do not have this trait measure for the backup participants.

<sup>10</sup> Restricting our sample to a young adult age range helped avoid additional confounding factors of age-related sleep changes. Depression risk was measured using the PHQ-2 (Kroenke et al, 2003) and anxiety risk using the

Once the subset of viable participants was identified from the sleep survey database, we then *ex ante* randomly assigned all viable participants to be recruited into the current study under the sleep restricted (SR: 5-6 hrs/night attempted sleep) or well-rested (WR: 8-9 hrs/night attempted sleep) prescribed sleep levels. These sleep levels are considered insufficient (SR) or sufficient (WR) based on guidelines of both the National Sleep Foundation and the U.S. Centers for Disease Control and Prevention (see <https://www.sleepfoundation.org/> and <https://www.cdc.gov/sleep/index.html> ) for the specific age range of the participants we studied. The invitation email sent to participants conveyed the sleep level assignment information by indicating that participant would be asked to spend between 5-6 or 8-9 hrs/night attempting sleep (participants only saw the sleep assignment level to which they were randomly assigned), and they would also wear a sleep tracking device and keep sleep diaries during the week. Thus, the sleep manipulation was not highly controlled as in the case of a sleep lab study, but participants knew we would use the objective sleep data from the tracking devices to verify good-faith efforts to comply prior to issuing the compensation tied to their sleep level compliance.

Recruited participants came to the lab for two distinct 1.5 hour sessions set one week apart on a Tuesday, Wednesday, or Thursday (to avoid weekend sleep effects) and at non-extreme times of day (between 10am and 6pm) to avoid significant circadian influences in the design. However, the exact time frame in which to attempt the prescribed sleep level was left up to the participant. Session 1 was used to complete Informed Consent, obtain some basic participant sleepiness, mood, and cognitive style measures (described later), and assign a sleep diary and wrist-worn accelerometer (sleep tracker or “actigraphy” device) to each participant. Paper and pencil sleep diaries common to sleep research were used, and these asked a few short questions both at bed time (lights out) and wake time from the participants. The sleep diaries were returned to the experimenters at the end of the study. We used research-grade devices that have been validated against polysomnographic measures of sleep time and, in conjunction with the complementary sleep diaries, the scored data are considered valid and objective regarding sleep levels (see Goldman et al, 2007 on details of the scoring protocol). Participants were

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GAD-7 instrument (Spitzer et al, 2006). Morningness-eveningness was assessed using the reduced form Horne and Östberg survey instrument (Adan and Almira, 1991; Horne, J. A., & Östberg, 1976), and we also assessed daytime sleepiness during the pre-screen online sleep survey using the Epworth daytime sleepiness measure (Johns, 1991). Finally, as an individual-specific, albeit subjective, characteristic, we also elicited the participant’s perceived optimal level of nightly sleep.

instructed on their use, but they had no need to worry about battery life or data downloads and simply carried on their daily routines except for the manipulation of their sleep levels. After completion of Session 1, the participant went home and did not return to the lab until 7 days later. Regular email reminders from the experimenter were sent to help keep in contact with participants and share necessary study information (e.g., reminders regarding protocol expectations, adherence even on weekends, cautionary notes if sleepy, upcoming Session 2, etc.).

The basic protocol for sleep manipulation was set up as a between-subjects version of the cross-over design protocol described in Dickinson et al. (2017)—the reduced statistical power of the between-subjects design is made up for with our larger sample size. This previous study helped inform our expectations regarding subject attrition and compliance with the prescribed sleep levels. Cohorts of up to 20 participants were recruited at a time for the 1-week experiment and each cohort included a mix of participants randomly assigned to the WR and SR treatments, which was private information. In total,  $n=279$  participants were recruited into the main study,  $n=258$  showed up for the initial Session, and  $n=233$  finished the full week protocol.

After the full week sleep manipulation, participants arrived at Session 2 in an experimentally-assigned SR or WR condition and were then administered the bargaining task after providing a reassessment of some self-reported sleepiness and mood metrics that were also elicited during Session 1—we discuss these measures later regarding tests of protocol validity. Here, participants within each cohort (ranging from 10-20 participants) were randomly matched with another participant for the 10-rounds. The matching remained in place for the entire 10-round bargaining task and so our design was a partners-matching protocol. In this case, the bargaining history of the pair may be particularly important in understanding current round behavior, which we controlled for in the data analysis. Bargaining was anonymous and via computer terminals (i.e., *not* face-to-face). Each cohort was made aware during Session 1 that not all participants in the cohort were prescribed the same sleep level, but it was also noted that the prescribed sleep levels were intended to be private information. As such, for the bargaining task the participants were blinded to the sleep treatment of one's counterpart, which is more similar to field environments where sleep status of others is typically not common knowledge.

The 10-round bargaining experiment involved zero-sum bargaining over the value of a variable,  $x$ , where the randomly assigned Player A's payoff decreased in the level of  $x$ , while the Player B's payoff increased in  $x$ . Even though pure zero-sum environments may not be the

norm, we view this stylized environment as a clean and controlled way to study bargaining and dispute whenever interests are misaligned to some degree. There is also evidence that individuals may view many environments as zero-sum even when they are not (Rubin, 2003), and this applies to organizations and workplace conflict situations as well (Carnevale and Pruitt, 1992). In our environment Player A submits offers,  $x_A$ , and Player B submits offers,  $x_B$ , to determine the value of  $x$ . The payoff functions (in dollars) used for each round were: Player A,  $\pi_A = 1.00 + (.005 \cdot (500 - x))$ ; Player B,  $\pi_B = 1.00 + (.005 \cdot (x - 500))$ . While  $x=500$  would “split the pie” and yield each a \$1.00 payoff for that round, the suggested bargaining range for Player A,  $x \in [200,700]$ , was different from that of Player B ( $x \in [300,800]$ ) and was not common knowledge. This design feature, based on Ashenfelter et al. (1992), is intended to make a simple 50-50 split of the pie less focal. The 10-rounds of bargaining were divided into two treatments, and the players could submit (but were not required to) up to five settlement proposals in any given round. There were no requirements that offers be made or that they improve upon previous offers, and so the offer stages provided only minimal structure to the environment. In the first treatment, five rounds of No Arbitration (NA) bargaining were administered. Here, impasse resulted in zero payoff for each bargainer. In these NA rounds, either Player A could accept Player B’s offer,  $x = x_B$ , or Player B could accept Player A’s offer,  $x = x_A$ , unless the offers overlapped,  $x_B \leq x_A$ , at which point a settlement at the average offer,  $x = \frac{x_A + x_B}{2}$ , was implemented.

After the NA treatment, five additional rounds of bargaining ensued under a Final-Offer Arbitration (FOA) procedure that mechanized the role of the arbitrator. Voluntary settlement would result if a bargainer accepted the other’s offer during any of the proposal stages for that round, or if the bargainers’ offers converged or overlapped, i.e.,  $x_B \leq x_A$ . Under the FOA rules, if settlement had not been reached after the five stages of proposal submissions, each bargainer was then required to submit a final offer. Assuming these final offers did not converge, in which case a last chance settlement resulted, the final offers were then compared to a draw,  $z$ , from the computerized arbitrator preference distribution,  $f(z)$ . The final offer closest to this arbitrator preference,  $z$ , was then implemented as the settlement value of  $x$  for that round.<sup>11</sup> Such

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<sup>11</sup> As noted in Ashenfelter et al. (1992), the use of a draw from a distribution of settlement preferences is a way to capture the real-world uncertainty disputants face regarding the likely preferences of an arbitrator (or, in our case, uncertainty regarding typical conflict outcomes or co-worker perceptions in such a dispute). The distribution of

uncertainty modeled as FOA here could also be interpreted as uncertainty surrounding how a manager may resolve the conflict using an informal FOA approach, or as uncertainty regarding friends' or co-workers' average opinion regarding one's position in the interpersonal conflict. A perception that the average opinion of others is aligned with one's own position, rather than the other disputant's position, improves self-esteem and implies a higher payoff. This methodology of a mechanized FOA procedure allows us to incorporate a common conflict outcome uncertainty across all rounds of our FOA treatment, which enables a clean identification of sleep's impact on settlement likelihood and facilitates our interpretation of the data on bargainer beliefs (i.e., optimism) in FOA. Given our focus on analyzing how sleep restriction impacts settlement rates, arbitration was monetarily costless in our environment, as in Ashenfelter et al. (1992). Others, however, have examined dispute under costly arbitration or with costly delay (e.g., Charness, 2000; Deck and Farmer, 2007). The assumption of zero monetary cost to bargainers of using FOA is not unreasonable even in some instances of formal arbitration use by the employer to help resolve worker conflict, because it would *not* be the workers (bargainers) who pay that explicit cost.

Our design is one in which dispute in NA is inefficient because resources are lost when conflict goes unresolved, while this is not the case of FOA. We justify this design feature by noting that our focus is not on conflict rate *level* comparisons between NA and FOA, but rather we consider the conflict rate in NA to be a measure of conflict tendency for a pair, and as such should be controlled for in the analysis of FOA conflict and how it may be affected by sleep. The comparison of conflict or settlement rates in a high (and certain) monetary cost NA environment with one or more arbitration environments where unresolved conflict does not result in resources loss—rather, it introduces *uncertainty* costs—also has precedence in the literature. Thus, the comparison of NA versus FOA conflict rates with similar studies in the extant

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arbitrator settlement preferences used in our environment was  $f(z) \sim N(500, 60)$ . Also following the Ashenfelter et al. (1992) approach, bargainers were not told of the specifics of the  $f(z)$  distribution. Rather, they were shown a table of 100 draws from the distribution and left to make their own inferences regarding expected outcomes. This makes the environment more parallel with field settings and how bargainers form beliefs from viewing past arbitrator outcomes. In the event of unbiased expectations of  $f(z)$ , the predicted final offers for risk neutral bargainers would be  $x_A=425$  and  $x_B=575$ . The predicted final offers would likely be closer to  $x_A=450$  and  $x_B=550$  given the asymmetric bargaining ranges used to limit mechanical split-the-pie outcomes (see details in Dickinson, 2004). Earlier research examined theoretical final offers of bargainers under FOA rules (Farber, 1980; Brams and Merrill, 1983), with others offering extensions for considerations such as multiple bargaining issues (Wittman, 1986) and private information (Samuelson, 1991).

literature can serve as a type of validation check on the bargaining environment we implemented. Our primary interest is the comparison of conflict rates in FOA across pairs with different numbers of SR bargainers. The parameterization and methods of the environment were based on the Ashenfelter et al. (1992) design as utilized in Dickinson (2004, 2005) to include elicitation of bargainer beliefs at the beginning of each FOA round and a within-subjects design where all participants experienced both NA and FOA treatments. The elicitation of beliefs was incentivized; participants were compensated an additional 20 tokens if their belief in that round was "...within 5 units of what the computer typically draws for X." The description of the belief elicitation was made right below the Table of 100 past draws of x shown to participants in the treatment instructions, which provided information on the likely draws of X used in the computerized FOA procedure (see Appendix B for these instructions). Because our treatment order always administered NA first, there was an order effect common across bargaining pairs. When analyzing a pair's dependency on FOA, we are careful to address this by incorporating a pair's full history of conflict as a main effect control variable in estimating the likelihood of dispute in FOA.<sup>12</sup> Instructions to the bargaining environment are in Appendix B. The participants were administered two other interactive decision tasks during the same session as well as individual tasks that generated individual characteristic measures but, given its relative complexity, we always administered this bargaining task first and so we can be assured that there is no impact of the other tasks on the data generation we report.

### *Hypotheses*

Regarding our hypotheses, we anticipated that our general findings should first replicate established results in the literature. Thus, we expected higher rates of unresolved conflict (i.e., "dispute rates") with FOA compared to NA, average beliefs should reflect optimism (i.e., beliefs will be biased in a self-serving direction), and the divergence of beliefs was hypothesized to predict FOA dispute likelihood. The theoretical extension in Dickinson (2006) highlights the impact of optimism on widening the predicted Nash equilibrium final offer gap of the bargainers, which we assumed implies an increased likelihood of dispute as well. We also hypothesized that

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<sup>12</sup> Alternatively, the common treatment order can be thought of as having given all bargaining pairs a training period of high (certain) cost NA bargaining before engagement in FOA, as many laboratory studies may start with some trial periods. Our main results focused on the sleep impacts on FOA and so reducing the interpretation of NA outcomes to those of practice rounds is fine, but we feel these outcomes are nevertheless worth reporting.

past use of FOA would increase the present use of the settlement procedure (the so-called “narcotic effect”). We did not make an explicit hypothesis regarding the impact of previous NA disputes on impasse within the NA treatment. The reason is that there are two opposing effects at work in NA. On the one hand, a history of unresolved conflict in NA may be predicted to harden the bargainers’ positions such that present voluntary settlement may be even less likely. On the other hand, dispute in NA is costly and so past dispute may make current dispute less likely via reinforcement learning. We are able to observe the net effect in our data as an exploratory analysis but would otherwise require a different experimental design to disentangle these effects. It is reasonable, however, to hypothesize that a pair’s NA conflict history may contribute towards settlement tendencies in FOA because the procedure guaranteed an easy alternative to voluntary settlement. Our approach to modeling likelihood of voluntary settlement in the current round as a function of past outcomes (settlement versus dispute) follows the general method of Bolton and Katok (1998), which is supported by more formal tests we conducted and report in Appendix C.

The novel hypotheses we present are with respect to how SR will impact bargaining outcomes. As noted from the literature highlighted above, we hypothesized that SR bargainers would be more optimistic in FOA compared to WR bargainers, and this was predicted to reduce voluntary settlement. One may also hypothesize reduced voluntary settlement among SR bargainers with FOA due to the fact that allowing FOA to resolve conflict is cognitively less demanding than the process of good-faith bargaining to achieve voluntary settlement, and we generally predicted SR would decrease the use of more cognitively demanding decision processes. Finally, SR bargainers were predicted to become more dependent on (or, addicted to) the use of FOA for similar reasons. Reduced efforts to negotiate voluntary settlements, or a stronger attachment to the pair’s conflict history as a status quo in their interactions, would both be considered less cognitively demanding strategies than engaging in good faith efforts to find mutually agreeable outcomes in FOA. In other words, the more that FOA has already been used by a bargaining pair in the past, the more effort it likely takes to overcome that history and voluntarily settle the current conflict. We therefore predict the tendency towards failed settlement efforts (i.e., dispute) will be magnified by sleep restriction given that SR typically promotes less cognitively demanding or deliberative processes. We articulate these hypotheses below, combining all hypotheses focused on replication of past findings as Hypotheses 1 (1a, 1b,

1c and 1d). The SR-focused hypotheses then follow. All other findings we report would be considered exploratory in nature.

H1a: Bargainers will display optimistic beliefs in FOA.

H1b: Dispute rates (i.e., rates of unresolved conflict) will be higher in FOA compared to NA.

H1c: Dispute in FOA will be positively related to optimism (i.e., the divergence of bargainer beliefs).

H1d: The use of arbitration (FOA) in the current round will be more likely the more it has been used in previous rounds (i.e., narcotic effect of arbitration).

H2: Beliefs about the arbitrator settlement draw in FOA will be more optimistic – biased in a self-serving direction – for SR compared to WR bargainers.

H3: Impasse will be more likely in FOA with more SR bargainers in the pair.

H4: The narcotic effect of FOA will be more pronounced with more SR bargainers in the pair.

### ***Protocol Validity***

We examined the participant sample characteristics and assessed validity of the protocol before analysis of the bargaining data. Regarding balance tests of the treatment assignment, Table 1 shows reasonable balance based on observable characteristics of our participants. Only *Depression Risk* was found to be marginally higher in the SR subsample, although these were *subclinical* levels of risk due to the fact that those at significant risk were excluded from recruitment during the prescreen phase.

The validity of the protocol was established by examining both objective and subjective measures as seen in Figure 1 and Table 2. Figure 1 shows the kernel density estimates of (actigraphy measured) total sleep time for the two sleep treatments. Here, we can clearly see the difference in objectively-measured nightly sleep between the two groups. The overlap of the two distributions highlights that a portion of our participants did not necessarily fully comply with our prescribed sleep levels. If one were to set an arbitrary compliance standard such that we deemed some *noncompliant*, then the stringency of the standard would dictate our compliance rate. One compliance standard may be to deem SR participants noncompliant if they slept more than 375 min/night (based on objective actigraphy measures) and WR participant noncompliant if sleeping less than 405 min/night. Under this standard we would have n=212 (SR 99, WR 113)

compliant participant. A stronger compliance standard requiring SR to sleep no more than 360 min/night and WR no less than 420 min/night would yield n=190 (SR 93, WR 97) compliant participants. Such standards are not totally arbitrary, as Figure 1 shows that they are data driven and essentially remove the subset of data that contributes to the sleep quantity distribution overlap (i.e., subjects assigned to one treatment but who are less statistically distinguishable from those assigned the opposite treatment). In what follows, we generally conducted analysis based on the weak compliance standard and then discuss any differences found if using the full sample (intent-to-treat) or stronger compliance standard in the sensitivity analysis.

In addition to the objective measure of total (nightly) sleep time, Table 2 presents nonparametric (Mann Whitney) test results from examining differences in other measures of interest. *Personal Sleep Dep* is a variable that subtracts one's nightly actigraphy measured sleep from the participant's perceived optimal amount of nightly sleep reported in the online sleep survey prior to the protocol. As such, it is a hybrid measure of objective sleep and subjective sleep need. Because this "sleep need" differs by individual, this measure could be useful assuming self-reported optimal sleep levels are accurate and not biased.<sup>13</sup> The *Karolinska* self-report sleepiness measure elicited a 1-9 rating of in-the-moment subjective sleepiness (commonly used in sleep research: Åkerstedt and Gillberg, 1990), and affective state measures for *Irritability* and *Alertness* were derived from the Positive and Negative Affect Scale (PANAS: Watson et al, 1988). We utilized the full combined measures of the positive and negative affect subscales of the PANAS in the *Positive Affect* and *Negative Affect* variables. *SleepManipulate* was a self-report of how much less (negative values) or more (positive values) the participant felt that treatment week made him/her sleep compared to that participant's regular habits.

Each measure, other than actigraphy sleep, *Personal Sleep Dep*, and *SleepManipulate* was elicited during Session 1 (the *Pre-Treatment* measure in Table 2) as well as at the beginning and end of Session 2—the *Post-Treatment* measure is the average of the measure taken at the beginning and end of the decision Session 2 (i.e., before and after the decision making tasks). Such measures allowed a within-subject test of the sleep treatment effect on that variable. Table

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<sup>13</sup> Most participants reported an optimal sleep level of more than what they report typically getting, and so in this sense perceived sleep need seemed consistent with conventional wisdom of sleep health experts regarding average insufficient sleep levels of most young adults. Also, because perceived personal optimal sleep was elicited on the online pre-screen survey at a point in time *prior* to the recruitment into this study, it is impossible that optimal sleep reports were somehow endogenous to the treatment assignment.

2 shows that only *Positive Affect* showed a difference across SR/WR assigned during session 1 (*pre-manipulation*), but all measures showed clear treatment impacts in the expected direction. Based on tests of these multiple measures, can conclude that our manipulation was successful—SR assignment significantly decreased objectively measured sleep, led to less-than-usual perceived sleep in participants, increased self-reported sleepiness and irritability, and reduced alertness and positive affect in general.

## RESULTS

### *Sample Selection*

Sample selection and attrition must be examined in our protocol, and we addressed this with probit analysis of two indicator variables: *Showed Up* indicates whether the subject who signed up for the study showed up for Session 1 or not, while *Finished Protocol* indicates whether a participant who started the protocol finished or was lost to attrition during the treatment week (i.e., subject withdrew). It is also worth noting that withdrawal may have been influenced by messaging from the experimenter who reviewed emailed bed/wake times during the protocol week and contacted participants who appeared out of compliance. A reminder requesting compliance or assessing participant well-being may cause a participant to withdraw from the study rather than put in effort to comply. Such experimenter influence is not present with respect to the *Showed Up* indicator variable given the experimenter has more limited and generic communication with all recruited participants prior to Session 1.

Probit estimation results in models (1) and (2) of Table 3 showed that, other than *SR* predicting attrition and selection into the final sample ( $p < .01$ ) only a couple of variables were marginally significant in predicting the likelihood of showing up for Session 1 or dropping out mid-week. Specifically, older subjects and those with higher self-reported *Optimal Sleep* need were marginally less likely to finish the protocol ( $p < .10$ ), and those with higher *Epworth* sleepiness scores (a validated measure of more chronic daytime sleepiness) were marginally less likely to show up ( $p < .10$ ) and finish the protocol ( $p < .05$ ) conditional on signing up. It is always possible that unobserved factors may play a role in selection or compliance (or behavior) of our participants. As such, even though we took additional measures to obtain a relatively robust set of sleep characteristic control measures on our participants, conclusions are subject to the usual caveat regarding possible omitted variables.

Regarding treatment assignment, *SR* strongly predicted a lower likelihood of finishing the protocol, conditional on starting it (the marginal effect estimate is a 17% reduced likelihood of finishing in *SR* compared to *WR* participants). Though *SR* assignment significantly reduced the likelihood of finishing the protocol, we will later see that *WR* assignment significantly decreased the likelihood of compliance with the sleep prescription, conditioned on completing the protocol. To address the issue of overall attrition in our sensitivity analysis, model (3) of Table 3 estimated inclusion in our final sample using the entire set of subjects recruited into the study. We use predicted probabilities of inclusion in the final sample to create inverse probability weights (*IPW*) for the subsequent regressions that will account for likelihood the participant is included in our sample. This approach is similar to that used to account for participant loss due to follow-up in epidemiological research.

### ***Individual Bargainer Analysis—Belief Optimism***

We first examined bargainer beliefs (H1a) and whether or not sleep restriction (*SR*) increased optimism (H2), as we hypothesized. For this, we could only examine data from the Final-Offer Arbitration (*FOA*) treatment, which contained the incentivized elicitation of bargainer beliefs regarding the likely outcome of the computerized settlement procedure. On average, the trend was in the direction of optimistic beliefs (average Player A belief of 487.59 and Player B belief of 507.36), but we required the analysis to condition on multiple beliefs per bargainer, as well as the key variable of treatment assignment, *SR*. Table 4 shows results from a series of estimations of bargainer beliefs in the *FOA* treatment. The dependent variable was coded to measure not the belief itself, but rather the optimism (or pessimism) of the belief with respect to the midpoint of the pie ( $x=500$ ). In this way, we can pool Players A and B and simply examine the *optimism* of belief of all bargainers. Regressions reported include simple regressions as well alternate specification that included controls for cohort fixed effects, treatment round, and the following subject-specific characteristics of interest: *Female* (=1), *ProSocial* (=1, based on the Social Value Orientation measure in Murphy et al., 2011), *CRT score* (from the 6-item Cognitive Reflection Task: Primi et al., 2016), and *Epworth* (the chronic daytime sleepiness measure taken in Session 1 prior to the treatment manipulation: Johns, 1991). A final model was estimated (model 4) that corrected for selection into the final sample (study attrition) using the inverse probability weight (*IPW*) of selection derived from the selection model (3) in Table 3.

The models in Table 4 were estimated on the subset of bargainers who met the weak-compliance standard regarding the sleep prescription, but results were qualitatively similar if using the full intent-to-treat sample or the more restrictive strong-compliance sample with the binary *SR* indicator for sleep state, as seen in Figure 2 (Model 1) and Appendix Tables A1 and A2. The summary of our findings regarding beliefs is that we did not find general support for H1a, the test of which is whether each models' constant term was statistically greater than zero (marginal support,  $p < .10$ , for H1a was only found in model (1) of Table 4 where we did not control for fixed effects, treatment round, demographics or sample selection). In other words, there was little evidence that beliefs were statistically different from  $x=500$  once we include a more complete set of control variables in the analysis. Regarding how sleep impacted optimism, we found some evidence supporting Hypothesis H2 across models (1)-(4) when using the binary *SR* indicator to control for sleep. Here, the positive and significant coefficient estimates on *SR* show evidence that *SR* increased the level of bargainer optimism in models (1) and (2) of Table 4. The significance level became more marginal ( $p < .10$  rather than  $p < .05$ ) as additional controls were added (model (3)) and when using the sample selection correction (model (4)). Overall, we consider the set of results in Table 4, and the complementary results using different sample restrictions in Appendix Tables A1 and A2, as at least weak support for Hypothesis H2.

Because we also have data on alternative measures of an individual's sleep state, additional analysis was conducted to examine whether the H2 finding is robust to using two alternative measures of sleep state as independent variables: *Karolinska* sleepiness reports, and *Personal Sleep Dep* as both defined above in the **Protocol Validity** subsection. We present the summary of the robustness analysis as coefficient plots in Figure 2, and the evidence suggests support for hypothesis H2 is not robust to sleep control measures other than the strictly exogenous binary *SR* indicator. Because the Figure 2 coefficient plots indicate that the choice of sleep control matters regarding H2, it is worth highlighting key differences in the various independent sleep measures considered. The binary indicator, *SR*, is the indicator of randomly assigned sleep treatment, over which we had control as experimenters. Though seemingly more informative due to their continuous nature, the variables measuring *Karolinska* (subjective) sleepiness and *Personal Sleep Dep*, are not entirely objective measures of sleep. And perhaps most importantly, these two alternative measures contain variation that is not entirely due to the random sleep treatment assignment. As such, there are econometric reasons to prefer the strictly

exogenous binary measure in of sleep in our analysis. In any event, we present the set of results for transparency and only endeavor to present the reader with the relevant information regarding the likely trade-offs faced in choosing one measure over the other. We now turn to our primary analysis of the determinants of voluntary settlement versus impasse.

### ***Pair-level Analysis—Dispute likelihood and the Narcotic Effect***

Up to this point our analysis has focused on beliefs, which are hypothesized to predict settlement likelihood, but a main interest in our experimental design was to examine determinants of voluntary settlement versus unresolved conflict. Table 5 reports estimated predictors of the dichotomous variable of impasse we called *Dispute* = 1 for rounds that ended without a voluntary settlement. Estimates shown are from linear probability models that estimate the determinants of *Dispute*. We estimated simple specifications, as well as specifications that included additional control variables, for analysis of the NA and FOA treatments. Results are similar using probit estimation (available on request), and so we present the linear model for simplicity of interpretation. We present analysis separated by treatment. The comparison of dispute rates across treatments shows strong support for H1b—dispute rates are, on average, 15 percentage points higher in FOA compared to NA.<sup>14</sup> This replicates dispute rate differences in the prior experimental literature that compares relatively similarly NA and FOA treatments (e.g., Ashenfelter et al, 1992; Dickinson, 2004, 2005).

The simple specifications estimated in Table 5 (models (1) and (4)) include controls for the number of SR bargainers in the pair (0, 1, or 2), pairwise beliefs in the FOA treatment, and dispute history controls. Pairwise beliefs in FOA treatments are measured by *Belief Gap*, which is positive when Player A's belief regarding the likely draw,  $z$ , of the computerized arbitrator was lower than Player B's belief—this can be considered a type of pairwise optimism with respect to bargainer beliefs that has been previously shown to predict bargaining failure in this environment (Dickinson, 2006). *Total Dispute History* captures the cumulative number of rounds the pair has previously disputed in their interactions, which allows a formal test of the

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<sup>14</sup> Average dispute rates depended on the sample (intent-to-treat, weakly or strongly compliant pairs) but average dispute rates are 14%-15% in NA treatments compared to 29%-32% in FOA rounds. This difference was statistically significant using a two-sample proportions test, or using regression estimates (clustering errors by pair) with or without group-specific controls, and with or without *IPW* selection correction ( $p < .01$  in all instances. Results available on request).

FOA treatment hypothesis H1d. The other interaction variable, *Treatment Disp Hist \* # SR*, was intended to examine whether any impact of dispute history differed based on the sleep composition of the bargaining pair, which tested hypothesis H5 in the FOA models. That is, did the impact of dispute history on the likelihood of FOA use depend on the number of SR bargainers in the pair? The impact of these variables on behavior in the NA treatments remains exploratory. Recall, rounds 1-5 are the NA treatment for all bargaining pairs, and so the *Treatment Dispute History \* # SR* interaction variable captured the interaction between the number of SR bargainers in the pair and the history of dispute *just* in the current treatment (i.e., its values range from 0-4). Therefore, our specification is such that, in the FOA treatment, we are also capturing the impact of the dispute history the pair had experienced in the NA treatments prior to engaging in FOA bargaining rounds—this will add a pure *level* effect to the dispute history of the pair given that the NA treatment rounds are completed once FOA begins (so, the number NA dispute rounds is fixed at that point). We include additional results in Figures A2 and A3 (Appendix A) to highlight the fact that our key results below are qualitatively similar when using alternative estimation approaches or specifications to control for dispute history (though the statistical significance level varies).<sup>15</sup>

The other specifications in Table 5 included additional controls for cohort fixed effects, treatment round (1-5), the gender and pro-sociality composition of the pair, and the average daytime sleepiness measure and average Cognitive Reflection Task score of the pair. To address the issue of sample selection, we estimated a final model for each treatment (models (3) and (6) in Table 5) that used the average *IPW* of the two bargainers in a weighted regression that corrected for the probability of inclusion in the final sample, as discussed in the ***Sample Selection*** subsection above. The error terms were also clustered at the bargaining pair level for the analysis. We further note that our decision to specify the probability of failed settlement (i.e., dispute) as a function of past dispute history reflects a view that learning within the bargaining pair is based on learning from past outcomes, rather than a persistent round-effect learning trend.

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<sup>15</sup> Results in Figure A2 of Appendix A were from estimations that separated the past dispute history in the FOA treatment into the “NA dispute history” and “FOA dispute history” components, while results in Figure A3 estimated separate models for bargaining pairs higher or lower based on different measures of “sleepiness”. Results showed consistency with the pattern of results reported in the main text (Table 5 and Figure 5). Namely, bargaining pairs with more history of using FOA were more likely to use FOA in the present round, and this history-dependency effect appeared stronger for more “sleepy” bargaining pairs, no matter how it was measured (see sensitivity analysis in Figures 3 and 5).

This is the approach suggested in Bolton and Katok (1998), which we tested in our data (results shown in Appendix C). Our data support the specification choice we present in Table 5 over a specification that only considers a persistent learning trend across rounds. The Appendix C test results also support our Table 5 findings regarding the directional effect that past dispute history has on the likelihood of current-round settlement in both NA and FOA treatments.

We report several important findings from Table 5, which are summarized also in the coefficient plots in Figure 3. First, an exploratory finding was that a higher number of past disputes had a tendency to *decrease* the likelihood of current round dispute in NA ( $p < .05$  in models (2) and (3) of Table 5). Additionally, this dispute history effect in monetarily costly NA bargaining was not impacted by the sleep composition of the pair. Secondly, consistent with prior research that supports H1c, we found that pairwise optimism predicted dispute as seen in models (4)-(6) of Table 5 ( $p < .05$ ). Finally, we estimated that the use of FOA can be addictive in the sense that past disputes increased the likelihood of using FOA in the present. The estimated positive and statistically significant coefficient estimate on *Total Dispute History* indicates that past dispute history increased the likelihood of impasse in FOA ( $p < .05$  across all FOA models in Table 5, which supports H1d).<sup>16</sup>

Importantly, when testing hypothesis H4 we estimated a significant positive effect of the interaction *Treatment Dispute History \* # SR* ( $p < .01$  for models (4) and (6), but  $p < .05$  for model (5) in Table 5). In other words, the increased likelihood of unresolved conflict in the current FOA round that is associated with a greater past history of unresolved conflict was larger for pairs with more sleep restricted bargainers. Another way of viewing this result is to say that *SR* contributed to a pair's decreased tendency to voluntarily settle conflict as experience with unresolved conflict increased. This supports our hypothesis H4 regarding the impact of *SR* bargainers on the narcotic effect in the FOA environment. In fact, the magnitude of the interaction effect indicates that one extra *SR* bargainer in the pair roughly doubled the predicted FOA narcotic effect for a given dispute history.<sup>17</sup> These results were all robust to correction for sample selection (models (3) and (6)).

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<sup>16</sup> As already noted, this estimated coefficient also captured the effect of how dispute history from NA rounds increased the likelihood of impasse in FOA bargaining. The estimation of additional specifications that differentiated dispute history by treatment showed qualitatively similar results (see Appendix Figures A2 and A3).

<sup>17</sup> This was true not just in the linear probability model, but in a probit estimation we found that the positive marginal effect on the interaction term *Treatment Disp Hist \* # SR* was statistically significant and larger in

### *The Narcotic Effect—sensitivity analysis*

Because these results in support of H4 represent our key finding (i.e., that SR bargainers show an increased “narcotic effect” or dependency on letting conflict go unresolved), we conducted extensive sensitivity analysis. Figures 4 and 5 present the results of the sensitivity analysis, and the details of these estimations are found in Appendix A (Tables A2 and A3). The sensitivity analysis involved estimating similar specifications for data meeting different standards of sleep-compliance (weakly-compliant, strongly-compliant, and intent-to-treat full sample estimates are compared). Additionally, as with the *Belief Optimism* estimations from Table 4, we also estimated specifications that varied the measure used to control for sleep at the level of the bargaining pair. *Pair Level Sleepiness* captured the average level of Karolinska subjective sleepiness of the pair, with higher values indicating overall increased sleepiness of the bargaining pair. *Pair Level Personal SD* measured the pair’s average level of personalized sleep deprivation, which combined the objective and *continuous* actigraphy measure of total sleep time and the participant’s self-report sleep need.<sup>18</sup>

The graphical depiction of the sensitivity analysis is in the form of coefficient plots of the NA (Figure 4) and the FOA model results (Figure 5) that present point estimates and confidence intervals of the key coefficient estimates: *Total Dispute History*, *Treatment Disp Hist \* # SR*, and *Belief Gap* for the FOA analysis. Thus, the sensitivity analysis represents a comprehensive examination of our key finding on different sleep-treatment compliance standards and on using binary assignment, subjective, and objective (continuous) measures to describe the sleep state of the bargaining pair.

The sensitivity analysis regarding the NA treatment showed that results were similar across samples and across pair-level sleep measures. There was no significant interaction between the sleep characteristic of the bargaining pair and the dispute history in predicting dispute in the current round. Also, though the significance of the main-effect of dispute history

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magnitude than the positive and significant marginal effect of *Total Dispute History*. These results are available on request.

<sup>18</sup> Though this measure is subject to some bias due to the self-report nature of optimal nightly sleep, recall that this measure was elicited in the online sleep survey at an early point in time compared to this experiment and so it at least poses no risk of being endogenously determined by the randomly assigned sleep treatment. In fact, self-reported optimal nightly sleep did not differ by sleep treatment assignment ( $7.98 \pm 1.21$  hrs/night for those assigned to SR and  $7.93 \pm .90$  hrs/night for those assigned to the WR treatment).

depended on the particular model estimated in Figure 3, the tendency was that more previous disputes led to an estimated reduced likelihood of current round dispute (i.e., an increased likelihood of settlement) in the monetarily costly NA environment. Regarding the FOA treatment, we see that the effect of *Belief Gap* on increasing the likelihood of dispute first seen in Table 5 and Figure 3 is robust to this sensitivity analysis, which further supports hypothesis H1c. The effect of past dispute history (*Total Dispute History*) was uniformly positive, but its significance was sensitive to the sample we used (Intent-to-Treat, Weak or Strong compliance) and the choice of pairwise sleep measure. Estimates from the full Intent-to-Treat sample (square-symbol point estimates in Figure 5) provided the most precise estimates of the positive dependency-effect evidence in our data. Finally, the 3 panels on the right-hand side of Figure 5 show the pairwise sleep interaction effects with the history of FOA dispute. We can see here that, no matter which sample used, which modeling choice regarding the sleep metric, or whether we use the *IPW*-correction or not, the evidence is robust in showing that more sleepy or SR bargaining pairs were estimated to have a stronger narcotic effect—for a given history of previous dispute, sleepy bargainers were even less likely to voluntarily settle a present-round conflict compared to well-rested bargainers when unresolved conflict implied an uncertain outcome (modeled by the FOA settlement procedure).

## **DISCUSSION**

This study focused on the effect of sleep on interpersonal conflict, and our data suggest that commonly experienced levels of insufficient sleep can contribute to the inability or unwillingness to resolve interpersonal conflict. In the workplace, for example, the right managerial strategy that appreciates the interplay of sleep and conflict resolution can therefore take advantage of opportunities to improve interpersonal relationships, the organizational culture, and company profits.

Consistent with previous research, our results indicate that optimism with respect to one's likely outcome in the event of unresolved conflict has a negative impact on the likelihood of a voluntary settlement. Insufficient sleep may have an indirect effect on the likelihood of voluntary conflict resolution via its impact on promoting a self-serving perspective, though our findings were more marginal and less robust on this question. Optimism is not always bad, but it likely harms the household or workplace culture when it results from an over-focus on one's own

position and a lack of the ability to see things from others' perspectives (Neale and Bazerman, 1983). Our main result, however, is that SR impromotes an increased dependency on risky conflict outcomes. This result is robust and consistent with the hypothesis that SR promoted a less effortful and deliberative approach to conflict resolution. For example, it may be that having more substantial history of unresolved conflict makes failed settlment the status quo for the bargaining pair. As such, our result could be interpreted as showing that SR bargainers were more prone to a type of status-quo bias. Alternatively, more sleepy bargainers may simply be less willing to put in the effort to voluntarily settle disputes as conflict history grows.<sup>19</sup>

Other implications of our findings for operational settings and managerial practice should be noted. First, our results speak specifically to how efforts to improve sleep hygiene can help improve conflict resolution. In occupational settings, the question of sleep in organizations has been largely ignored (Barnes, 2011; Giurge, 2017; Barnes and Watson, 2019) and, to the extent it has been examined, any focus on how sleep may impact conflict resolution in the workplace is mostly absent in the literature. Workplace Health Promotion Programs (WHPPs) are already quite common in many organizations, but efforts to target sleep in WHPPs should be increased. Examples of such efforts have been suggested by others and include: sleep awareness or fatigue management training, the design of job scheduling (reducing long work hours or minimizing shift rotations), family-friendly policies (e.g., parental leave), promoting timed naps in the workplace, modifying workplace environmental characteristics (e.g., lighting), or referral for sleep disorder treatment (Christian and Ellis, 2011; Barnes, 2011; Redeker et al, 2019). Having a workplace culture that promotes healthy sleep behaviors, as opposed to promoting workaholic tendencies<sup>20</sup>, can help minimize levels of unrealistic expectations that may be at the root of workplace conflict.<sup>21</sup> When conflict occurs, our findings suggest that more rested workers are less likely to be caught in a bad habit of requiring others to help resolve their conflict (or letting

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<sup>19</sup> While it is true that both risk attitude and expectations contribute to one's acceptable settlement, past research suggests that expectations are more important, in general (Dickinson, 2009). Of course, given the present study does not directly measure risk attitude, an important question for future research would be to examine the relative impact on beliefs versus risk attitude of a sleepy versus a well-rested bargainer (e.g., risk aversion is suggested as an explanation for relatively high settlement rates in the use of formal FOA in Major League Baseball salary disputes (Hanany et al, 2007)). Both effects argue for the effective management of beliefs and perceived risk to reduce workplace conflict.

<sup>20</sup> As noted in Barnes (2011), the former CEO of SYNEX Canada, Jim Estill, is an example of someone who created a workplace culture that suggested sleep is for wimps and reflects lack of motivation.

<sup>21</sup> Babcock et al (1997) highlight the value of attempts to improve expectations and "de-bias" bargainers involved in conflict.

it go unresolved altogether). This is important given that voluntary settlement of conflict by the interested parties is generally preferred for maintaining a positive organizational culture. In workplace settings, these effects of sleep on conflict resolution may be mediated by one's commitment level, job satisfaction, and organizational citizenship behaviors. Importantly, each of these has been found to suffer when workers have lower sleep levels (Barnes et al, 2013). Similar factors may mediate how sleep impacts conflict resolution in other settings (e.g., friendship commitment, marital satisfaction, etc).

While a bargaining environment with explicit costs to third-party arbitration would have made it possible to estimate explicit welfare differences between sleepy and more rested bargaining pairs, we note that voluntary settlements are generally considered superior to arbitrated settlements (Crawford, 1979). As such, we feel it reasonable to suggest that a sleepy individuals will generate implicit welfare costs associated with the reduced voluntary resolution of conflict. Because our results indicate that individuals can become dependent or more accepting of uncertain-outcome alternatives to voluntary settlements (i.e., someone else will fix it, or let the conflict go unresolved), the fact that sleep health of spouses or workers can help increase the likelihood of voluntary settlement is worth emphasis. Overall, our findings can be made practical in at least a couple of ways: (1) Efforts to de-bias unrealistic or self-serving expectations can improve conflict management.<sup>22</sup> This is a general results in both our SR and WR participants, though we also have some marginal evidence that WR promotes a somewhat reduced levels of self-serving expectation; (2) Efforts to improve sleep hygiene help limit the general tendency to become addicted to the risky impasse outcome that results from unresolved conflict.

One criticism may be that our findings are drawn from an anonymous bargaining environment without face-to-face negotiations. One may speculate that sleepy bargainers in real-world environments may compensate via their ability to process nonverbal language or read emotion and facial cues, and thus successfully navigate conflict. However, existing research as well as an additional finding in our study suggest otherwise. Van Der Helm et al (2010) showed

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<sup>22</sup> In a workplace context, a hybrid mediator-arbitrator alternative style of management may also be effective (Ross and Conlon, 2000). Yet another managerial approach recently suggested would be to engage in hostile fashion with both workers involved in dispute, such that the manager becomes the common enemy in a way that can actually facilitate resolution of the conflict (Zhang et al, 2017). However, this approach is not without risk given that hostility towards only one or the other worker in conflict (or relatively more hostility towards one, perhaps) can eliminate this beneficial conflict management effect.

that total sleep deprivation impairs one's ability to accurately read emotion in facial expressions, which presents an additional challenge to expectations that sleepy bargainers may successfully navigate negotiations. As noted earlier, Kahn-Greene et al (2006) also reported that extreme total sleep deprivation harms one's ability to respond in a healthy way to frustrating interpersonal situations. Both of these studies involved extreme total sleep deprivation. Here, we offer a final result from our study that may speak to this issue at least with regards to moderately sleep restricted individuals.<sup>23</sup> Specifically, we administered half of the Eyes test of social intelligence (Baron-Cohen et al, 2001) to our participants during Session 1 and the other half during Session 2 of our study after the bargaining task, such that we have a measure of the treatment impact on one's ability to identify emotion states. We conducted estimates of the determinants of the change in the number of correct Eyes-image emotions detected pre- and post-manipulation (positive values indicate one identified *more* correct emotions post-manipulation). Table 6 shows our estimates as a function of three sleep metrics, similar to our approach in the sensitivity analysis of our main narcotic effect result: intent-to-treat (SR indicator), self-reported (Karolinska) sleepiness ratings, and *Average Nightly Sleep* night derived from the objective actigraphy measurements.<sup>24</sup> By all measures, we have a robust finding that our SR participants performed significantly worse on the Eyes test relative to baseline when compared to WR participants relative to baseline. As such, even in our sample of more moderate partial sleep restriction, this evidence would suggest that SR participants are unlikely to overcome negative behavioral impacts of their sleepiness by merely adding a face-to-face element to conflict resolution process. Though not the main focus of this paper, a reduced ability to read facial queues when sleepy would present yet another obstacle to successfully navigating conflict in the workplace (see Elfenbein et al, 2007, on the benefits of emotion recognition in a negotiations environment).

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<sup>23</sup> The Eyes test results were not mentioned among our main results because they had no direct bearing on the analysis of our anonymous bargaining environment that involved no face-to-face interactions. The Eyes test was administered in our design as a measure of simple social intelligence that may yet be relevant to our overall understanding of sleep impacts in negotiations environments.

<sup>24</sup> The sensitivity analysis in Figure 4 and 5 used *Personal Sleep Dep* as opposed to the *Total Sleep Time* measure which incorporates one's perceived optimal sleep time. If using *Personal Sleep Dep* to predict the change in Eyes test scores, we found a qualitatively similar result (i.e., more *Personal Sleep Dep* predicts a decrease in % Eyes correct post treatment), but the precision of the estimate was low ( $p = .147$ ).

One may still highlight the limitation of using a stylized bargaining environment to make claims regarding negotiation settings of practical importance (e.g., workplace disputes, relational disputes). Here, we have studied zero-sum bargaining as a starting point for these efforts, and zero-sum environments are relevant in some but not all instances of interpersonal conflict. Natural extensions of this research would be to examine how insufficient sleep impacts bargaining in positive-sum games, when bargainer power differentials exist, or when longer term interactions are present. While not as directly applicable to the conflict environment we study, other research has highlighted the links between sleep and workplace bullying and note, in particular, the likely causal impact of workplace conflict on sleep problems (see the meta analysis in Mielsen et al, 2020).<sup>25</sup> Our design precludes the possibility that conflict in our task impacts the sleep measures we have on our participants, but an additional benefit to an organization of healthy sleep exists if improved conflict resolution helps reduce the negative feedback that conflict can have on sleep health going forward (see also Gordon et al, 2017, who recognize the likely bidirectional impact of sleep on social process regulation).

Because our design implemented a particular set of FOA rules to resolve dispute, one may also wonder whether our findings generalize to modeling the resolution of uncertainty differently or to different information conditions (e.g., conventional arbitration rules do not constrain the final arbitrated settlement as in FOA, or full information environments rather than an incomplete information environment where counterpart payoffs are not fully known). Clearly, one can model the uncertainty of unresolved conflict in multiple ways. The novel results presented here can be viewed as one part of a hopefully more comprehensive examination of the issue of sleep and bargaining. This is not meant to be the last word on the topic, but rather a first glimpse into the impact of sleep on bargaining that may help stimulate additional research. The likely impact of insufficient sleep on interpersonal inter-personal conflict or formal bargaining in personal and occupational settings is, as our results suggest, more widespread than previously understood.

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<sup>25</sup>Fox and Stallworth (2009) noted the cost to an employer of defending an unlawful discrimination lawsuit, which can be considered a rather high-stakes type of workplace conflict, may be close to \$100,000. Such prospects make more salient the workplace benefits of efforts and help promote voluntary settlement of a variety of types of workplace conflict, such as through promotion of mediation over arbitration. Promoting healthy sleep would be an indirect approach to help promote voluntary resolution of workplace conflict.

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**TABLE 1:** Balance Test of treatment assignment on individual characteristics

| <b>Variable</b>  | <b>SR (n=103)</b><br>Mean (st. error) | <b>WR (n=130)</b><br>Mean (st. error) | <b>T-stat</b><br>(2-sided p-value) |
|------------------|---------------------------------------|---------------------------------------|------------------------------------|
| Female (=1)      | .553 (.049)                           | .592 (.043)                           | 0.593 ( $p = .553$ )               |
| Race             | 5.010 (.049)                          | 4.900 (.080)                          | -1.169 ( $p = .243$ )              |
| Ethnicity        | 1.922 (.027)                          | 1.923 (.023)                          | 0.021 ( $p = .983$ )               |
| Age              | 19.311 (.295)                         | 19.685 (.245)                         | 0.975 ( $p = .331$ )               |
| Optimal Sleep    | 7.981 (.092)                          | 8.077 (.090)                          | 0.752 ( $p = .453$ )               |
| Anxiety Risk     | 4.058 (.261)                          | 4.162 (.237)                          | 0.293 ( $p = .770$ )               |
| Depression Risk  | .801 (.083)                           | .631 (.068)                           | -1.630 ( $p = .105$ )              |
| Epworth          | 7.864 (.338)                          | 8.169 (.335)                          | 0.641 ( $p = .522$ )               |
| Reduced-MEQ      | 13.136 (.281)                         | 12.885 (.279)                         | -0.635 ( $p = .526$ )              |
| 6-item CRT score | 2.117 (.169)                          | 2.369 (.151)                          | 1.118 ( $p = .265$ )               |

**Notes:** Both Depression and Anxiety risk are sub-clinical levels of risk with means representing scores on from the PHQ-2 and GAD-5 questionnaires, respectively, administered during the pre-screen sleep and demographic survey. Those at significant risk of either were not recruited as this was an exclusion criterion.

**TABLE 2:** Protocol Validation on sleep and mood measures

| Variable                                   | Pre-treatment (Session 1)            |                                      |                              | Post-treatment (Session 2)           |                                      |                              |
|--|--------------------------------------|--------------------------------------|------------------------------|--------------------------------------|--------------------------------------|------------------------------|
|  | SR<br>(n=103)<br>Mean<br>(std. dev.) | WR<br>(n=130)<br>Mean<br>(std. dev.) | Difference<br>(Mann-Whitney) | SR<br>(n=103)<br>Mean<br>(std. dev.) | WR<br>(n=130)<br>Mean<br>(std. dev.) | Difference<br>(Mann-Whitney) |
| <b>Nightly Sleep<br/>(min/night)</b>       | ---                                  | ---                                  | ---                          | 321.96<br>(32.01)                    | 439.37<br>(34.07)                    | $z = 12.699^{***}$           |
| <b>Personal Sleep Dep<br/>(min/night)</b>  | ---                                  | ---                                  | ---                          | 156.87<br>(57.25)                    | 44.82<br>(67.63)                     | $z = -10.566^{***}$          |
| <b>Sleep Manipulate<br/>∈ [-4, +4]</b>     | ---                                  | ---                                  | ---                          | -2.66<br>(1.31)                      | 1.14<br>(1.58)                       | $z = 12.009^{***}$           |
| <b>Karolinska<br/>(sleepiness ∈ [1,9])</b> | 4.88<br>(1.71)                       | 4.60<br>(1.73)                       | $z = -1.12$                  | 6.38<br>(1.55)                       | 3.77<br>(1.42)                       | $z = -10.070^{***}$          |
| <b>Irritable<br/>∈ [1, 5]</b>              | 1.53<br>(0.74)                       | 1.54<br>(0.78)                       | $z = -0.16$                  | 2.42<br>(1.08)                       | 1.68<br>(.84)                        | $z = -5.597^{***}$           |
| <b>Alert<br/>∈ [1, 5]</b>                  | 3.29<br>(0.77)                       | 3.27<br>(0.88)                       | $z = -.132$                  | 2.34<br>(.96)                        | 3.48<br>(.73)                        | $z = 8.536^{***}$            |
| <b>Positive Affect<br/>∈ [10, 50]</b>      | 31.57<br>(5.95)                      | 33.01<br>(6.22)                      | $z = 1.73^*$                 | 23.28<br>(7.72)                      | 32.09<br>(7.18)                      | $z = 7.854^{***}$            |
| <b>Negative Affect<br/>∈ [10, 50]</b>      | 14.05<br>(3.69)                      | 14.28<br>(3.79)                      | $z = 0.605$                  | 15.68<br>(4.79)                      | 14.46<br>(4.43)                      | $z = -2.257^{**}$            |

**Notes:** \* $p < .10$ , \*\* $p < .05$ , \*\*\* $p < .01$  for the 2-tailed test.  $n = 232$  (SR 103, WR 129) for Mann-Whitney tests on nightly sleep due to one WR participants for whom complete actigraphy sleep data were not available. A regression of the *change* in Positive Affect from pre- to post-treatment indicates the difference is significant and we can conclude that the SR condition significantly decreased Positive Affect. In general, the treatment impact of assignment to SR was greater than that of WR assignment on the measures obtained both pre-treatment and post-treatment (see Appendix Figure A1 for more details).

**TABLE 3:** Sample Selection and Attrition Analysis

| <b>Probit Estimation</b> | <b>Dep Var =<br/>Show Up (=1)</b><br><i>(Conditional on recruited)</i> | <b>Dep Var =<br/>Finished Protocol (=1)</b><br><i>(Conditional on Session 1 Show up)</i> | <b>Dep Var =<br/>Finished Protocol (=1)</b><br><i>(Conditional on recruited)</i> |
|--------------------------|--|--|--|
| <u>Variable</u>          | (1)<br>Coefficient (SE)  | (2)<br>Coefficient (SE)  | (3)<br>Coefficient (SE)  |
| SR (=1)                  | 0.323 (0.240)  | -1.690 (0.411)***  | -.568 (.192)***  |
| Female (=1)              | -0.190 (0.261)   | -0.221 (-.272)   | -.156 (.203)   |
| Minority (=1)            | 0.054 (0.371)  | -0.252 (0.320)   | .096 (.248)  |
| Age                      | -0.009 (0.044)   | -0.067 (0.036)*  | -.047 (.030)   |
| Optimal Sleep            | 0.033 (0.118)  | -0.246 (0.131)*  | -.117 (.095)   |
| Anxiety Risk             | -.007 (0.049)  | -0.077 (0.054)   | -.035 (.039)   |
| Depression Risk          | 0.127 (0.173)  | 0.181 (0.169)  | .138 (.131)  |
| Epworth                  | 0.073 (0.039)*   | 0.043 (0.039)  | .062 (.030)**  |
| Reduced-MEQ              | 0.048 (0.039)  | 0.024 (0.042)  | .039 (.031)  |
| Observations             | N=279  | N=258  | N=279  |
| Log Likelihood           | -69.440  | -61.958  | -115.01616   |

**Notes:** Full recruited sample of n=279 participants, n=258 started the protocol (i.e., showed up for Session 1) and n=233 finished the protocol. \* $p < .10$ , \*\* $p < .05$ , \*\*\* $p < .01$  for the 2-tailed test. Model (3) used to determine weights for selection correction based on inverse probability weighting in individual outcomes analysis.

**TABLE 4:** Predictors of *Belief Optimism* (weakly compliant participants)

| Dependent Variable= <i>Belief Optimism</i> (measured as (500- <i>Belief</i> ) for Players A and ( <i>Belief</i> -500) for Players B such that positive values indicate optimistic beliefs independent of the bargainer's role. |                  |                  |                  |                  |
|--|------------------|------------------|------------------|------------------|
| Variable   | (1)<br>Coef (SE) | (2)<br>Coef (SE) | (3)<br>Coef (SE) | (4)<br>Coef (SE) |
| Constant   | 4.821 (3.540)*   | 3.537 (10.678)   | 7.417 (15.465)   | 2.610 (2.040)    |
| SR   | 9.300 (5.181)**  | 9.635 (5.537)**  | 8.732 (5.67)*    | 8.704 (5.625)*   |
| Cohort Fixed Effects   | No               | Yes              | Yes              | Yes              |
| Total Dispute History  | ---              | ---              | 2.308 (1.986)    | 2.610 (2.040)    |
| Treatment Round  | ---              | ---              | -1.391 (1.222)   | -1.527 (1.353)   |
| Female   | ---              | ---              | -9.080 (5.647)   | -8.927 (5.680)   |
| ProSocial  | ---              | ---              | 2.537 (6.387)    | 2.213 (6.067)    |
| CRT-6  | ---              | ---              | -.537 (1.671)    | -.487 (1.631)    |
| Epworth  | ---              | ---              | -.565 (1.671)    | .359 (.659)      |
| <i>IPW Selection Correction</i>  | No               | No               | No               | Yes              |
| Observations   | 1060             | 1060             | 1055             | 1055             |
| subjects   | 212              | 212              | 211              | 211              |
| $R^2$ (overall)  | .006             | .025             | .034             | .032             |

**Notes:** \* $p < .10$ , \*\*  $p < .05$ , \*\*\* $p < .01$  for the 2-tailed test. Hypothesized effects are examined using the appropriate 1-tailed test (i.e., the test on the estimated constant term for Hypothesis 1a and the test of the coefficient on the SR indicator variable for Hypothesis 2). Estimations are on the subsample of data meeting the weak compliance standard (SR slept no more than 375 min/night, WR slept no less than 405 min/night). Models (1)-(3) are random effect GLS regressions. The selection-corrected models in (4) is a linear regression estimated with robust standard errors clustered on subject. Selection corrections are based on the inverse probability weights (*IPW*) from Table 3 model (3). See comparable results in Appendix A for estimations on the Intent-to-treat (full) sample and the Strongly compliant sample (Tables A1 and A1). The Social Value Orientation (SVO) measure was not captured for one participant during Session 1, which reduced the sample by 5 observations when including controls (models (3) and (4)).

**TABLE 5: Predictors of Dispute. Weakly compliant bargaining pairs.**

Random effects GLS regressions

| Dep Var = Dispute (0,1)<br>Variable    | NA Treatment<br>(Avg dispute rate = 14.3%) |                   |                   | FOA Treatment<br>(Avg dispute rate = 29.8%) |                     |                     |
|--|--|-------------------|-------------------|---|---------------------|---------------------|
|  | (1)<br>Coef (SE)                           | (2)<br>Coef (SE)  | (3)<br>Coef (SE)  | (4)<br>Coef (SE)                            | (5)<br>Coef (SE)    | (6)<br>Coef (SE)    |
| Constant                               | .178<br>(.039)***                          | .359<br>(.120)*** | .333<br>(.121)*** | .156<br>(.047)***                           | -.310<br>(.221)     | -.261<br>(.214)     |
| Belief Gap                             | ---  | ---               | ---               | .0005<br>(.0002)**                          | .0005<br>(.0002)*** | .0006<br>(.0002)*** |
| Total Dispute History (0-9)            | -.104<br>(.062)*                           | -.131<br>(.052)** | -.128<br>(.051)** | .087<br>(.026)***                           | .056<br>(.028)**    | .066<br>(.028)**    |
| # SR in pair (0,1,2)                   | -.012<br>(.029)                            | -.013<br>(.034)   | -.010<br>(.033)   | -.011<br>(.036)                             | .041<br>(.040)      | .029<br>(.039)      |
| Treatment Disp Hist * # SR             | .044<br>(.062)                             | .033<br>(.055)    | .028<br>(.053)    | .072<br>(.023)***                           | .061<br>(.030)**    | .080<br>(.031)***   |
| Cohort Fixed Effects                   | No   | Yes               | Yes               | No  | Yes                 | Yes                 |
| IPW Selection Correction               | No   | No                | Yes               | No  | No                  | Yes                 |
| Treatment Round                        | ---  | -.021<br>(.012)*  | -.018<br>(.012)   | ---   | .017<br>(.013)      | .007<br>(.013)      |
| # Females in pair (0,1,2)              | ---  | -.038<br>(.025)   | -.038<br>(.025)   | ---   | .054<br>(.041)      | .049<br>(.039)      |
| # ProSocial in pair (0,1,2)            | ---  | .017<br>(.033)    | .023<br>(.032)    | ---   | .0009<br>(.040)     | -.0001<br>(.038)    |
| Avg Epworth daytime sleepiness in pair | ---  | -.013<br>(.007)*  | -.013<br>(.007)*  | ---   | -.001<br>(.009)     | -.0006<br>(.008)    |
| Avg CRT score in pair                  | ---  | .027<br>(.016)*   | .029<br>(.016)*   | ---   | .087<br>(.020)***   | .079<br>(.019)***   |
| n                                      | 460  | 460               | 460               | 460   | 460                 | 460                 |
| pairs                                  | 92   | 92                | 92                | 92  | 92                  | 92                  |
| Wald $X^2$                             | 3.84                                       | 90.48             | ---               | 110.11                                      | 239.82              | ---                 |
| $R^2$ (overall)                        | .0122                                      | .0906             | .0872             | .1914                                       | .2459               | .2500               |

\* $p < .10$ , \*\*  $p < .05$ , \*\*\* $p < .01$  for the 2-tailed test. Random effects GLS regressions used for models (1), (2), (4), (5). One-tailed tests are appropriate for our directional hypotheses and so certain significance levels identified above are conservative. Results are similar with random effects Probit estimations. Sample selection corrected models (3) and (6) are ordinary least squares estimates. Selection-corrected models (3) and (6) use an average level of the inverse probability weight of the two bargainers, with probability weights based on Table 3 (model 3) predictions. All models (1)-(6) cluster errors at the bargaining pair level. Estimations are conditioned on both bargainers in pair being weakly compliant. See Appendix Tables 2A and 2B for estimates on full Intent-to-Treat sample and subsample of strongly compliant pairs.

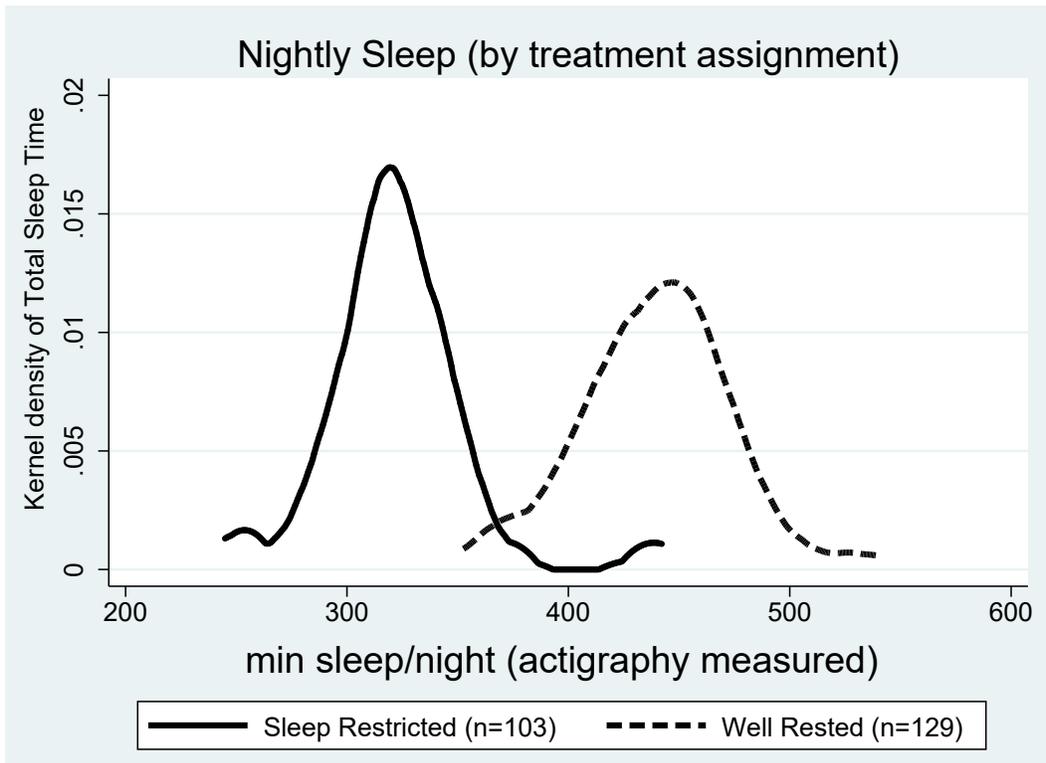
**TABLE 6: Eyes Test scores post versus pre-treatment**

**Dep Var = Change in % Eyes Correct (Post-Pre treatment)**

| <b>Variable</b>       | <b>(1)</b><br>Coef (SE) | <b>(2)</b><br>Coef (SE) | <b>(3)</b><br>Coef (SE) | <b>(4)</b><br>Coef (SE) |
|-----------------------|-------------------------|-------------------------|-------------------------|-------------------------|
| constant              | 0.009<br>(0.012)        | -0.115<br>(0.115)*      | -0.261<br>(0.082)***    | -0.101<br>(0.069)       |
| SR (=1)               | -0.040<br>(0.019)**     | -0.047<br>(0.018)**     | ---                     | ---                     |
| Avg Nightly Sleep     | ---                     | ---                     | 0.0003<br>(0.00014)**   | ---                     |
| Karolinska Sleepiness | ---                     | ---                     | ---                     | -0.011<br>(0.004)**     |
| Female (=1)           | ---                     | 0.036<br>(0.019)*       | 0.030<br>(0.019)        | 0.042<br>(0.019)**      |
| Minority (=1)         | ---                     | -0.018<br>(0.022)       | -0.012<br>(0.022)       | -0.018<br>(0.022)       |
| Age                   | ---                     | 0.007<br>(0.003)**      | 0.006<br>(0.003)**      | 0.007<br>(0.003)**      |
| CRT-score             | ---                     | -0.005<br>(0.005)       | -0.004 (0.005)          | -0.004<br>(0.005)       |
| observations          | 233                     | 233                     | 233                     | 233                     |
| R-Squared             | .019                    | 0.063                   | 0.057                   | 0.062                   |

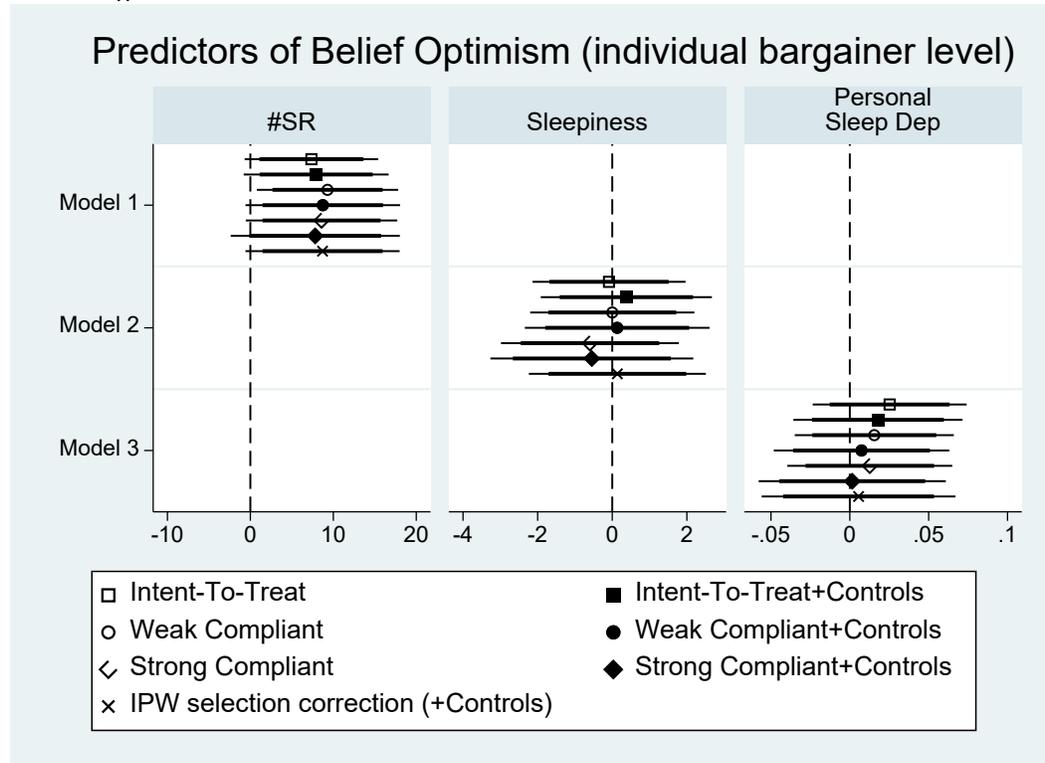
**Notes:** \* $p < .10$ , \*\* $p < .05$ , \*\*\* $p < .01$  for the 2-tailed test. Results are qualitatively similar if using our constructed measure of “Personal Sleep Deprivation” as the sleep measures (i.e., higher levels of *Personal SD* decrease the *Change in % Eyes Correct*), although the effect for this variable is not as precisely measured ( $p = .147$ ). Recall the *Personal SD* measure is constructed from the actigraphy measured nightly total sleep time subtracted from the individual’s self-assessed optimal nightly sleep time prior to the experiment (i.e., in the online pre-screen survey).

**FIGURE 1:** Average nightly sleep by treatment assignment



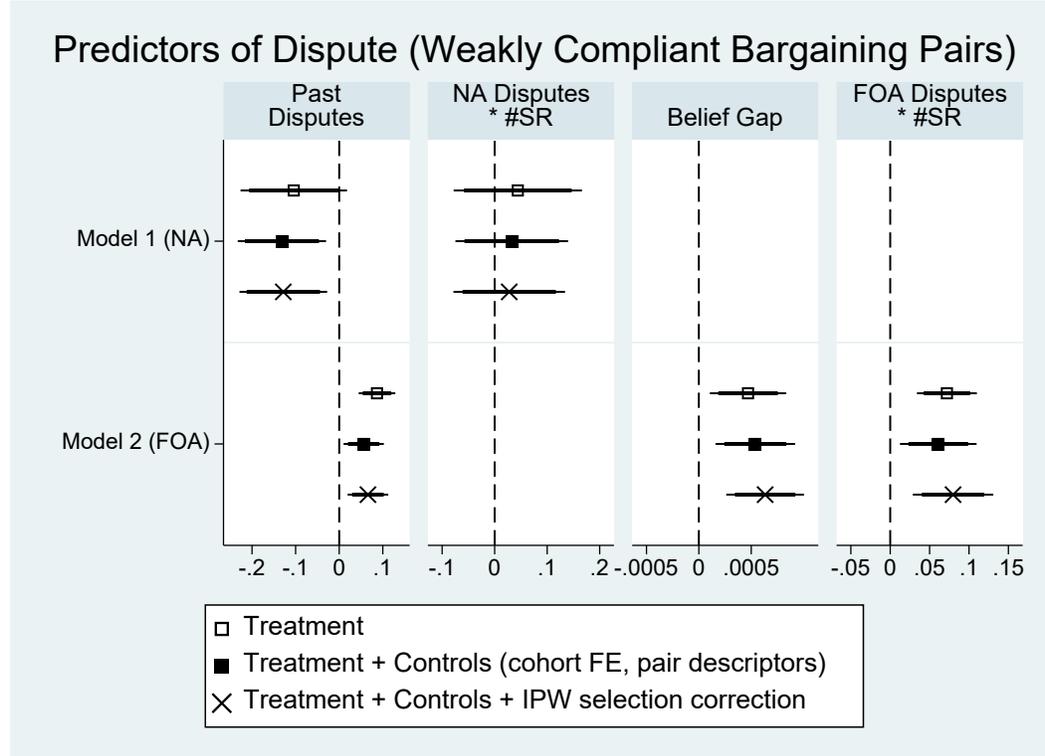
**Note:** Actigraphy data were incomplete for one of the participants completing the protocol, reducing the sample by one participant when examining actigraphy data.

**FIGURE 2:** Estimates of the impact of sleep impact on *Belief Optimism* under different sample restrictions, different specifications, and using different independent variable measures for sleep effect. (see Table 4 and Appendix Tables A1 and A1 for more detailed estimation results of models using the SR indicator))



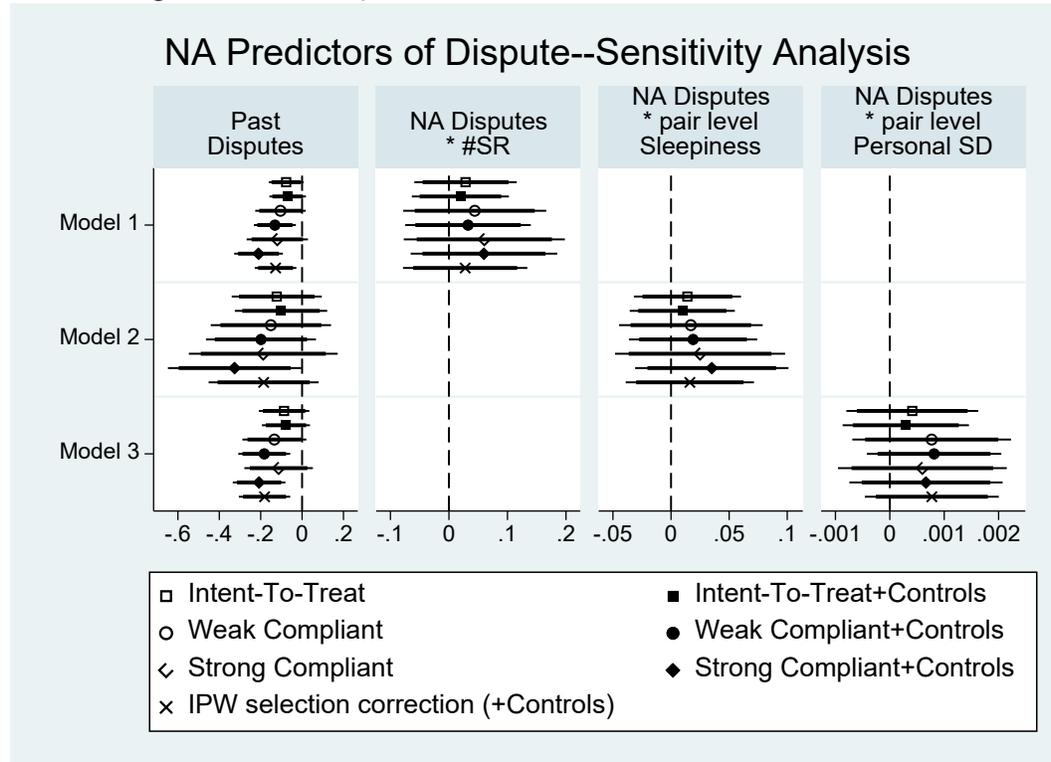
**Notes:** Coefficient name at top of the column. Controls for the sleep attribute of the bargainer are as follows: *Model 1*—SR indicator  $\in [0,1,2]$ ; *Model 2*—the *Karolinska* self-report sleepiness measure  $\in [1,9]$ ; *Model 3*—the *Personal Sleep Deprivation* level (in minutes per night)  $\in [-180, 288]$ . Controls included the cumulative number of previously disputed rounds for that bargainer, cohort fixed effects, round controls for gender, prosociality, daytime sleepiness, and the CRT score of that individual. Point estimates are shown with their 90% (thicker line) and 95% (thinner line) confidence intervals assuming a 1-tailed test for the directional hypothesis H2.

**FIGURE 3:** Estimates of Narcotic Effect and Impact of # SR in pair on the narcotic effect (by treatment) (see Table 5 for detailed estimation results)



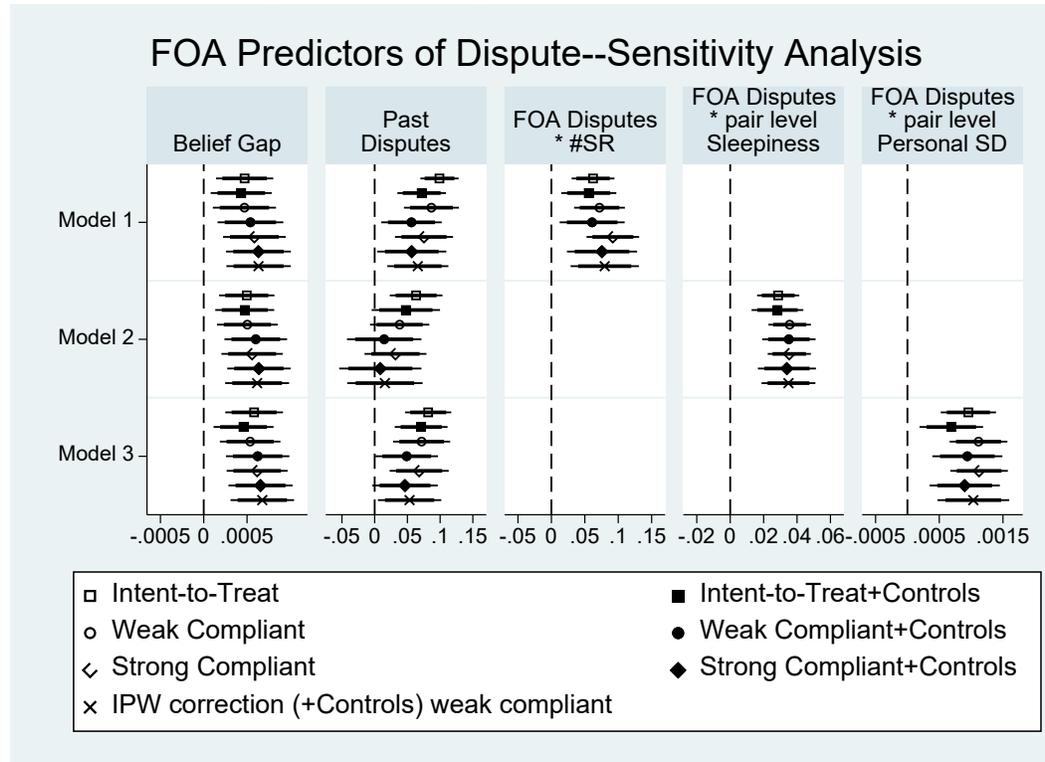
**Notes:** “Past Disputes” measures the impact of cumulative dispute history on currently round likelihood of dispute. Interaction effect measures the additional impact of the number of SR bargainers in the pair on this dependency effect within the treatment. Sensitivity analysis shows coefficient estimates for intent-to-treat, weakly and strongly compliant pairs, as well as models with and without pair specific fixed effects and controls. Squares identify coefficient estimates, and bars represent the 95% (thinner line) and 90% (thicker line) confidence intervals for the 2-tailed test for Model 1 (NA) but the 1-tailed tests for Model 2 (FOA) given our directional hypotheses regarding the impact of past disputes (H1d), beliefs (H2), and SR effects on dispute dependency in FOA (H5).

**FIGURE 4:** Estimates of the Likelihood of Dispute (=1). Robustness of *null* “narcotic” effect in **No Arbitration** Treatment (see Table 5 and Appendix Tables A3 and A4 for detailed estimation results of models using the SR indicator)



**Notes:** Coefficient name at top of the column. All models controlled for the gap in Beliefs (positive Gap = optimistically divergent beliefs). Controls for the sleep attribute of the pair are as follows: *Model 1*—the number of SR bargainers in the pair  $\in [0,1,2]$ ; *Model 2*—the average level of sleepiness of the two bargainers  $\in [1,9]$ ; *Model 3*—the average level of personal sleep deprivation (in minutes per night) of the two bargainers  $\in [-60, 275]$ . Controls included cohort fixed effects, round controls, pairwise controls for the number of prosocial and females in the pair, the average level of daytime sleepiness, and the average CRT scores of the pair. Point estimates are shown with their 90% (thicker line) and 95% (thinner line) confidence intervals assuming a 2-tailed test.

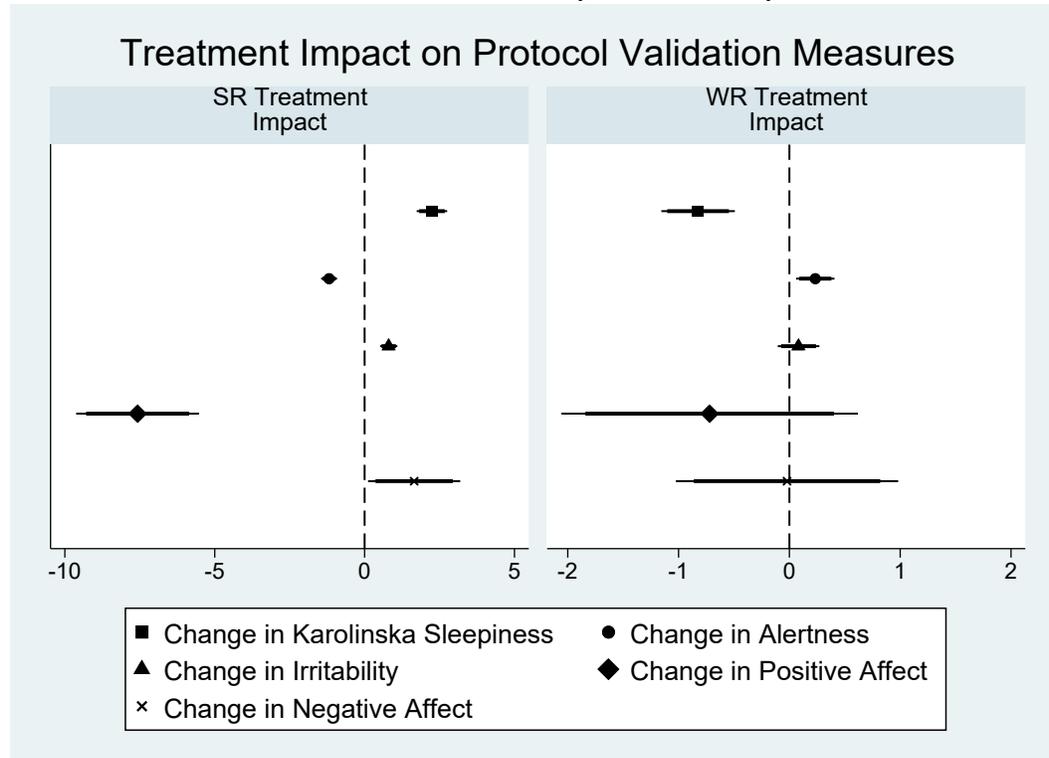
**FIGURE 5:** Estimates of the Likelihood of Dispute (=1). Robustness of Belief and Narcotic Effect estimates in **Final Offer Arbitration** (see Table 5 and Appendix Tables A3 and A4 for detailed estimation results of models using the SR indicator)



**Notes:** Coefficient name at top of the column. All models controlled for the gap in Beliefs (positive Gap = optimistically divergent beliefs). Controls for the sleep attribute of the pair are as follows: *Model 1*—the number of SR bargainers in the pair  $\in [0,1,2]$ ; *Model 2*—the average level of sleepiness of the two bargainers  $\in [1,9]$ ; *Model 3*—the average level of personal sleep deprivation (in minutes per night) of the two bargainers  $\in [-60, 275]$ . Controls included cohort fixed effects, round controls, pairwise controls for the number of prosocial and females in the pair, the average level of daytime sleepiness, and the average CRT scores of the pair. Point estimates are shown with their 90% (thicker line) and 95% (thinner line) confidence intervals assuming a 1-tailed test appropriate for our directional hypotheses for the variables listed above.

## APPENDIX A: Additional Estimations (sensitivity analysis)

FIGURE A1: Protocol Validation--FULL SAMPLE (Intent-to-Treat)



**Notes:** Scale of x-axis (measuring change in characteristics) differs in the left and right panels. Coefficient plot results are from simple regressions of the pre-treatment to post-treatment *change* in each measure on a constant term and an indicator for SR. Thus, the change in the measure for those assigned to the WR treatment is captured by the constant term, and the coefficient on the SR indicator measure the impact of assignment to the SR treatment on the change in the characteristic of interest. The thin (thick) lines represent the 95% (90%) confidence intervals on the estimate (2-tailed tests, though 1-tailed would be appropriate for the tests of SR assignment on Karolinska sleepiness, Alertness, and Irritability).

**TABLES A1, A2: See Figure 2 in main manuscript for coefficient plots of the key coefficient estimates of these Tables along with those from Table 4 (the *Weakly Compliant*) sample estimates and additional estimations that replace the SR indicator with alternative measures of sleep.**

**TABLE A1: Predictors of *Belief Optimism*--FULL SAMPLE (Intent-to-Treat)**

| Dependent Variable= <i>Belief Optimism</i> (measured as (500- <i>Belief</i> ) for Players A and ( <i>Belief</i> -500) for Players B such that positive values indicate optimistic beliefs independent of the bargainer's role. |                  |                  |                  |                  |
|--|------------------|------------------|------------------|------------------|
| Variable   | (1)<br>Coef (SE) | (2)<br>Coef (SE) | (3)<br>Coef (SE) | (4)<br>Coef (SE) |
| Constant   | 6.751 (3.194)**  | 2.056 (9.854)    | 5.671 (15.005)   | 6.231 (12.873)   |
| SR   | 7.363 (4.895)*   | 7.912 (5.131)*   | 7.932 (5.312)*   | 7.981 (5.178)*   |
| Cohort Fixed Effects   | No               | Yes              | Yes              | Yes              |
| Total Dispute History  | ---              | ---              | 1.709 (1.827)    | 2.386 (1.865)    |
| Treatment Round  | ---              | ---              | -.928 (1.165)    | -1.077 (1.311)   |
| Female   | ---              | ---              | -7.660 (5.312)   | -7.464 (5.147)   |
| ProSocial  | ---              | ---              | 4.813 (6.202)    | 4.523 (5.756)    |
| CRT-6  | ---              | ---              | -1.055 (1.580)   | -1.044 (1.633)   |
| Epworth  | ---              | ---              | .457 (.727)      | .533 (.627)      |
| <i>IPW Selection Correction</i>  | No               | No               | No               | Yes              |
| Observations   | 1210             | 1210             | 1160             | 1160             |
| subjects   | 242              | 242              | 232              | 232              |
| $R^2$ (overall)  | .004             | .021             | .030             | .029             |

**Notes:** \* $p < .10$ , \*\* $p < .05$ , \*\*\* $p < .01$  for the 2-tailed test. Hypothesized effects are examined using the appropriate 1-tailed test (i.e., the test on the estimated constant term for Hypothesis 1a and the test of the coefficient on the SR indicator variable for Hypothesis 2). Estimations are on the full sample of participants. Models (1)-(3) are random effect GLS regressions. The selection-corrected models in (4) is a linear regression estimated with robust standard errors clustered on subject. Selection corrections are based on the inverse probability weights (*IPW*) from Table 3 model (3). Reduced observations in columns (3) and (4) models reflects fact that one or more characteristic controls measures was not completed by 10 participants.

**TABLE A2: Predictors of *Belief Optimism*-- Strongly compliant bargaining pairs.**

| Dependent Variable= <i>Belief Optimism</i> (measured as (500- <i>Belief</i> ) for Players A and ( <i>Belief</i> -500) for Players B such that positive values indicate optimistic beliefs independent of the bargainer's role. |                  |                  |                  |                  |
|--|------------------|------------------|------------------|------------------|
| Variable   | (1)<br>Coef (SE) | (2)<br>Coef (SE) | (3)<br>Coef (SE) | (4)<br>Coef (SE) |
| Constant   | 5.512 (3.967)    | 4.027 (10.934)   | 8.640 (16.945)   | 8.907 (15.049)   |
| SR   | 8.579 (5.551)*   | 8.777 (6.012)*   | 7.724 (6.196)    | 7.934 (6.016)*   |
| Cohort Fixed Effects   | No               | Yes              | Yes              | Yes              |
| Total Dispute History  | ---              | ---              | 2.291 (2.140)    | 3.027 (2.216)    |
| Treatment Round  | ---              | ---              | -.915 (1.310)    | -1.133 (1.444)   |
| Female   | ---              | ---              | -0.869 (6.120)   | -9.683 (6.076)   |
| ProSocial  | ---              | ---              | .979 (7.202)     | .778 (6.918)     |
| CRT-6  | ---              | ---              | -1.321 (1.874)   | -1.162 (1.854)   |
| Epworth  | ---              | ---              | .414 (.819)      | .542 (.686)      |
| <i>IPW Selection Correction</i>  | No               | No               | No               | Yes              |
| Observations   | 950              | 950              | 945              | 945              |
| subjects   | 190              | 190              | 189              | 189              |
| $R^2$ (overall)  | .005             | .026             | .037             | .035             |

**Notes:** \* $p < .10$ , \*\*  $p < .05$ , \*\*\* $p < .01$  for the 2-tailed test. Hypothesized effects are examined using the appropriate 1-tailed test (i.e., the test on the estimated constant term for Hypothesis 1a and the test of the coefficient on the SR indicator variable for Hypothesis 2). Estimations are on the subsample of data meeting the strong compliance standard (SR slept no more than 360 min/night, WR slept no less than 420 min/night). Models (1)-(3) are random effect GLS regressions. The selection-corrected models in (4) is a linear regression estimated with robust standard errors clustered on subject. Selection corrections are based on the inverse probability weights (*IPW*) from Table 3 model (3). Data on one participant (5 rounds) was unavailable for the estimations with controls in models (3) and (4) models due to the fact that a characteristic control measure was not completed by the participant meeting the strong compliance sample condition.

**TABLES A3, A4: See Figures 4 and 5 in main manuscript for coefficient plots of the key coefficient estimates of these Tables along with those from Table 5 (the *Weakly Compliant*) sample estimates and additional estimations that replace the SR indicator with alternative measures of sleep.**

**TABLE A3: Predictors of Dependency (Addiction) to Dispute. FULL SAMPLE (Intent-to-Treat)**  
Random effects GLS regressions that include *cumulative* dispute history variable.

| Dep Var = Dispute (0,1)                | NA Treatment<br>(Avg dispute rate = 14.7%) |                    |                    | FOA Treatment<br>(Avg dispute rate = 29.5%) |                    |                    |
|--|--|--------------------|--------------------|---|--------------------|--------------------|
|  | (1)<br>Coef (SE)                           | (2)<br>Coef (SE)   | (3)<br>Coef (SE)   | (4)<br>Coef (SE)                            | (5)<br>Coef (SE)   | (6)<br>Coef (SE)   |
| Constant                               | .168<br>(.030)***                          | .379<br>(.123)***  | .352<br>(.121)***  | .152<br>(.036)***                           | -.139<br>(.192)    | -.114<br>(.183)    |
| Belief Gap                             | ---  | ---                | ---                | .0005<br>(.0002)*                           | .0004<br>(.0002)** | .0005<br>(.0002)** |
| Total Dispute History (0-9)            | -.077<br>(.043)*                           | -.070<br>(.045)    | -.067<br>(.045)    | .099<br>(.018)***                           | .072<br>(.023)***  | .084<br>(.022)***  |
| # SR in pair (0,1,2)                   | -.001<br>(.024)                            | .006<br>(.027)     | .009<br>(.027)     | -.021<br>(.030)                             | .014<br>(.036)     | .003<br>(.035)     |
| Treatment Disp Hist * # SR             | .028<br>(.045)                             | .020<br>(.042)     | .016<br>(.041)     | .062<br>(.019)***                           | .056<br>(.025)**   | .074<br>(.026)***  |
| Cohort Fixed Effects                   | No   | Yes                | Yes                | No  | Yes                | Yes                |
| IPW Selection Correction               | No   | No                 | Yes                | No  | No                 | Yes                |
| Treatment Round                        | ---  | -.034<br>(.011)*** | -.031<br>(.012)*** | ---   | .012<br>(.013)     | .002<br>(.013)     |
| # Females in pair (0,1,2)              | ---  | -.036<br>(.021)*   | -.035<br>(.021)*   | ---   | .031<br>(.032)     | .029<br>(.030)     |
| # ProSocial in pair (0,1,2)            | ---  | -.004<br>(.025)    | .001<br>(.025)     | ---   | -.008<br>(.035)    | -.008<br>(.033)    |
| Avg Epworth daytime sleepiness in pair | ---  | -.005<br>(.006)    | -.006<br>(.006)    | ---   | -.0008<br>(.008)   | .00003<br>(.008)   |
| Avg CRT score in pair                  | ---  | .031<br>(.013)**   | .032<br>(.013)**   | ---   | .070<br>(.020)***  | .062<br>(.019)***  |
| n                                      | 605  | 555                | 555                | 605   | 555                | 555                |
| pairs                                  | 121  | 111                | 111                | 121   | 111                | 111                |
| Wald $\chi^2$                          | 4.32                                       | 51.55              | ---                | 164.65                                      | 208.43             | ---                |
| $R^2$ (overall)                        | .009                                       | .070               | .066               | .199  | .226               | .231               |

\* $p < .10$ , \*\*  $p < .05$ , \*\*\* $p < .01$  for the 1-tailed test on hypothesized effects in the FOA treatment models regarding *Belief Gap*, *Total Dispute History*, *#SR in pair*, and *Treatment Dispute History \* #SR*. Otherwise, two-tailed test values reported. Random effects GLS regressions used for models (1), (2), (4), (5). Results are similar with random effects Probit estimations. Sample selection corrected models (3) and (6) are ordinary least squares estimates. Selection-corrected models (3) and (6) use an average level of the inverse probability weight of the two bargainers, with probability weights based on Table 3 (model 3) predictions. All pairs included in analysis.

**TABLE A4:** Predictors of Dependency (Addiction) to Dispute. **Strongly compliant bargaining pairs.** Random effects GLS regressions that include *cumulative* dispute history variable.

| Dep Var = Dispute (0,1)                   | NA Treatment<br>(Avg dispute rate = 15.3%) |                    |                    | FOA Treatment<br>(Avg dispute rate = 31.9%) |                     |                     |
|---|--|--------------------|--------------------|---|---------------------|---------------------|
|   | (1)<br>Coef (SE)                           | (2)<br>Coef (SE)   | (3)<br>Coef (SE)   | (4)<br>Coef (SE)                            | (5)<br>Coef (SE)    | (6)<br>Coef (SE)    |
| Constant                                  | .214<br>(.050)***                          | .523<br>(.142)***  | .506<br>(.144)***  | .185<br>(.062)***                           | -.172<br>(.248)     | -.157<br>(.247)     |
| Belief Gap                                | ---  | ---                | ---                | .0006<br>(.0002)***                         | .0006<br>(.0002)*** | .0007<br>(.0002)*** |
| Total Dispute History<br>(0-9)            | -.120<br>(.076)                            | -.210<br>(.060)*** | -.211<br>(.060)*** | .075<br>(.027)***                           | .057<br>(.032)**    | .061<br>(.032)**    |
| # SR in pair (0,1,2)                      | -.035<br>(.035)                            | -.073<br>(.039)*   | -.075<br>(.040)*   | -.028<br>(.032)                             | .015<br>(.049)      | .017<br>(.049)      |
| Treatment Disp Hist<br>* # SR             | .060<br>(.070)                             | .060<br>(.064)     | .057<br>(.063)     | .092<br>(.024)***                           | .076<br>(.032)***   | .082<br>(.032)***   |
| Cohort Fixed Effects                      | No   | Yes                | Yes                | No  | Yes                 | Yes                 |
| <i>IPW Selection<br/>Correction</i>       | No   | No                 | Yes                | No  | No                  | Yes                 |
| Treatment Round                           | ---  | -.008<br>(.014)    | -.006<br>(.014)    | ---   | .009<br>(.015)      | .005<br>(.015)      |
| # Females in pair<br>(0,1,2)              | ---  | -.066<br>(.031)**  | -.066<br>(.031)**  | ---   | .060<br>(.036)*     | .054<br>(.036)      |
| # ProSocial in pair<br>(0,1,2)            | ---  | .011<br>(.038)     | .016<br>(.038)     | ---   | .015<br>(.045)      | .012<br>(.044)      |
| Avg Epworth daytime<br>sleepiness in pair | ---  | -.021<br>(.008)**  | -.021<br>(.008)**  | ---   | -.008<br>(.011)     | -.007<br>(.011)     |
| Avg CRT score in pair                     | ---  | .034<br>(.019)*    | .035<br>(.019)*    | ---   | .063<br>(.021)***   | .059<br>(.020)***   |
| n   | 385  | 385                | 385                | 385   | 385                 | 385                 |
| pairs                                     | 77   | 77                 | 77                 | 77  | 77                  | 77                  |
| Wald $X^2$                                | 3.17                                       | 79.74              | ---                | 126.45                                      | 308.98              | ---                 |
| $R^2$ (overall)                           | .013                                       | .123               | .120               | .194  | .251                | .256                |

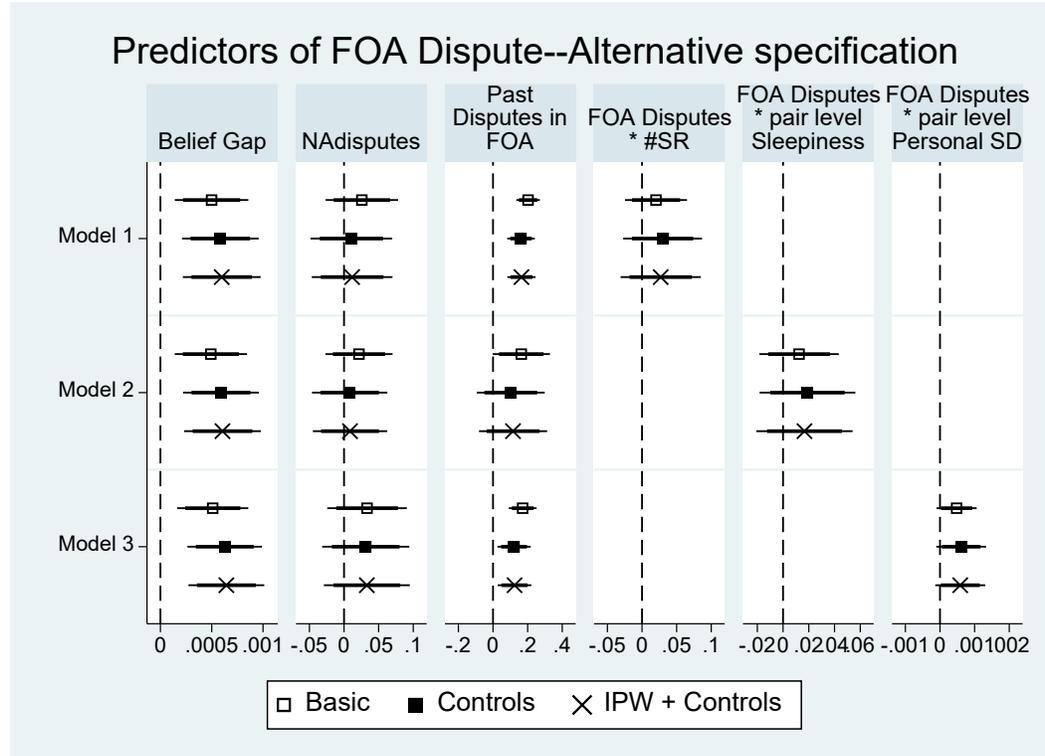
\* $p < .10$ , \*\*  $p < .05$ , \*\*\* $p < .01$  for the 1-tailed test on hypothesized effects in the FOA treatment models regarding *Belief Gap*, *Total Dispute History*, *#SR in pair*, and *Treatment Dispute History \* #SR*. Otherwise, two-tailed test values reported. Random effects GLS regressions used for models (1), (2), (4), (5).

Results are similar with random effects Probit estimations. Sample selection corrected models (3) and (6) are ordinary least squares estimates. Selection-corrected models (3) and (6) use an average level of the inverse probability weight of the two bargainers, with probability weights based on Table 3 (model 3) predictions. Estimations are conditioned on both bargainers in pair being strongly compliant (SR no more than 360/night and WR no less than 420 min/night sleep).

### Additional Sensitivity Analysis of FOA Narcotic Effect (compare with main text Figure 5)

An alternative specification to that shown in Table 5 and Figure 5 (narcotic effect analysis of FOA treatment) would be to separate the controls for history of past dispute in NA and the history of dispute in FOA. The models shown in Table 5 combined this into a singular *Total Dispute History* variable. As such, the interaction term *Treatment Dispute History \* numSR* interacted the number of SR bargainers in the pair with only the variable portion of this history once in the FOA treatment—recall that NA always precedes FOA treatment and so the total number of disputes in NA was a fixed part of the *Total Dispute History* variable. Figure A2 below shows results from estimations where the specification controlled separately for previous history of dispute in the NA treatment (*NAdisputes*), which was fixed for the pair in the analysis of FOA treatment rounds, from the previous history of disputes in FOA. One can see from these estimations conducted using weakly compliant pairs subset of data, that the narcotic effect (i.e., the interaction term in each model) remained consistently positive but no longer precisely estimated. Similar results were found using the subset of Strongly Compliant pairs or the full sample (Intent-to-Treat), with key interactions terms always positive and ranging from  $p = .046$  to  $p = .272$  (results available on request).

**FIGURE A2:** Estimates of the Likelihood of Dispute (=1). Robustness of Belief and Narcotic Effect estimates in **Final Offer Arbitration—Alternative Specification (weakly compliant pairs)**

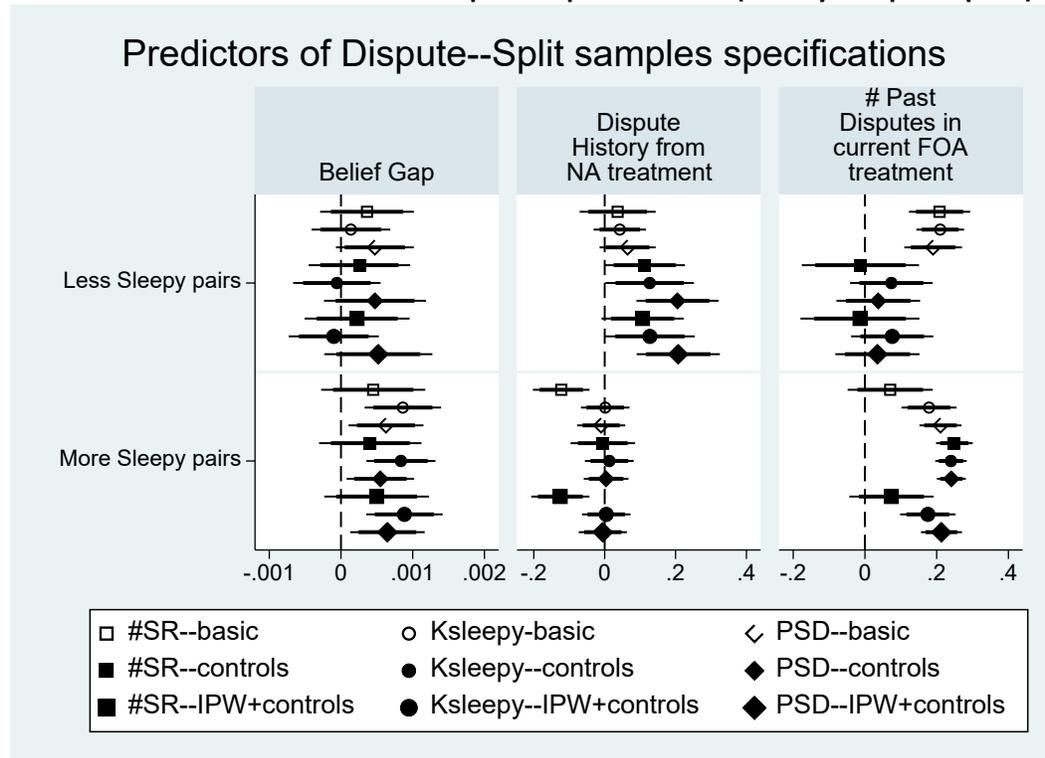


**Notes:** Coefficient name at top of the column. Controls for the sleep attribute of the pair are as follows: Model 1—the number of SR bargainers in the pair; Model 2—the average level of sleepiness of the two bargainers; Model 3—the average level of personal sleep deprivation of the two bargainers. Controls included cohort fixed effects, round controls, pairwise controls for the number of prosocial and females in the pair, the average level of daytimes sleepiness, and the average CRT scores of the pair. Point estimates are shown with their 90% (thicker line) and 95% (thinner line) confidence intervals for the appropriate 1-tailed test given our prior hypotheses.

Because the precision of the narcotic effect estimation appeared somewhat sensitive to the specification used to control for previous NA history, we present below another alternative approach to the estimation of the narcotic effect. Specifically, we conducted a series of split-sample analysis where separate models were estimated for the subset of bargaining pairs either above or below median level of the 3 pair-specific measures of sleep state from the main analysis: *numSR*, pair level sleepiness (*Ksleepy*) and pair level Personal SD (*PSD*). The total number of weakly compliant bargaining pairs used for analysis was the full set of 92 weakly compliant pairs for the *PSD* sample split, 84 pairs for the *Ksleepy* sample split (i.e., eliminating all pairs for which average sleepiness was at the median level *Ksleepy*=5), but only 53 pairs for the *numSR* sample split (eliminating all *numSR*=1 pairs and comparing only the *numSR*=0 pairs (n=27) with the *numSR*=2 pairs (n=26)).

As can be seen in Figure A3 below, the more sleepy (by all measures) bargaining pairs displayed a fairly robust level of addiction to FOA (increased likelihood of FOA use the more it was used in the past by that pair) compared to the level estimated for less sleepy pairs. The only measure for which this was not a statistically significant finding was the *numSR* sample split, which suffered from reduced sample size and statistical power (even here, the estimated narcotic effect was significant using the *Ksleepy* measure to capture pairwise sleep). These results were similar if estimated on the full intent-to-treat sample or the subset of strongly compliant bargaining pairs (available on request).

**FIGURE A3:** Estimates of the Likelihood of Dispute (=1). Robustness of Belief and Narcotic Effect estimates in **Final Offer Arbitration—Split sample estimates (weakly compliant pairs)**



**Notes:** Coefficient name at top of the column. Basic models includes only covariates shown in each panel. Models with controls include cohort fixed effects, round, and pair-specific controls as in Figure 5 (main text) specifications. IPWs+controls specifications also correct for selection bias using the average inverse probability weights for the pair from the selection equation (Table 3). Less (More) Sleepy pairs are those below (above) median on pair-specific measures (#SR in pair, average level of Karolinska sleepiness in pair, average level of Personal Sleep Deprivation in pair). Point estimates are shown with their 90% (thicker line) and 95% (thinner line) confidence intervals for the appropriate 1-tailed test given our prior hypotheses.

## APPENDIX B: Bargaining Experiment Instructions

*Notes:* several decision tasks were presented to all participants during the decision task Session 2 of the protocol, referred to below as “today’s experiment”. Specifically, we first captured self-report measures of participant sleepiness, treatment impact, and mood. Then, the “six stages” described below were as follows: Stage 1=bargaining task, Stage 2=Centipede Game, Stage 3=matrix games, Stage 4=Social Value Orientation task, Stage 5=Race game, Stage 6=Eyes test. The bargaining task was always administered first in this set of stages. Earnings for each incentivized “stage” were withheld until the end of the experiment so that earnings from any one stage would not create a wealth effect that may impact behavior in subsequent stages. Baseline sleepiness, mood, and other measures used as individual-specific controls in our analysis were obtained during Session 1 prior to the sleep manipulation.

Again, each component (task or survey measure) was referred to as a “stage” of the overall experiment. The bargaining task was Stage 1 of this decision session (hence, the reference to the “first stage” in the instructions below, which is referring to the entire 2-treatment 10-round bargaining experiment). Instructions for Player A shown. Differences between instructions presented to Player B are noted in square brackets and shaded. The experiment was administered over computerized interface.

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### GENERAL INSTRUCTIONS:

#### OVERVIEW

Today's experiment is divided into six stages. Every stage is, in all respects, separate from the others. This means that the choices you make in each stage will have no effect on your earnings in the other stages. Your final earnings will be the sum of the earnings obtained in the six stages. At the beginning of each stage, we will provide you with the relevant instructions and give examples to ensure that the experiment is clear.

In each stage of the experiment you will earn tokens. At the end of the experiment, tokens will be converted to cash. Each token is worth \$0.10 or 10 tokens is worth \$1.00. At the end of the experiment, your earnings will be issued based on an ID number, not based on your name. The payments will be issued in cash in an envelope. This way none of the other participants will know how much you earned.

#### Bargaining Experiment (first stage) PLAYER A [PLAYER B]

The first stage will consist of 10 decision rounds. In this stage, you have been randomly matched with an anonymous counterpart and you will remain matched with the same person for all 10 decision rounds of this stage. You have also been randomly assigned the role of Player A [Player B] in this stage, and your counterpart is assigned the role of Player B [Player A].

For each of the decision rounds of this stage, you and your counterpart will be given a fixed number of times (5 opportunities) to mutually agree upon the value of a variable, X. Your range of possible X values lies from 200 to 700 [300 to 800] in increments of one (this may not be the same range as that for your counterpart). The value of X at the end of the round will determine your earnings for that round. As a Player A [Player B], your earnings in any given round are larger for lower [higher] values of X (and your counterpart, Player B [Player A], earns more for higher [lower] values of X). More specifically, you would earn 250 tokens in a decision round if the value of X is 200 [800], while you would earn 0 tokens in that round for X=700 [300]. If you agree to a value of X above the maximum [below the minimum] of your suggested bargaining range (i.e., X greater than 700 [lower than 300]) then you would earn a negative payoff. Each increase [decrease] in the value of X of 1 unit lowers your payoff by ½ token. During the computerized bargaining, a tool on the screen will tell you what your payoff would be for different X values. Each new round will have you and your counterpart bargain again of the value of X, and so this process repeats in each round of the stage.

You will only interact with your anonymous counterpart via the computer screen, and during each bargaining round, you will be allowed to propose up to 5 values of X (i.e., your offer to your counterpart of what you would like X to be). Similarly, your counterpart can also propose his/her desired value of X up to 5 times. At any point during the bargaining round, you are allowed to end the round by "accepting" the current offer of your counterpart, and your counterpart may also end the round by accepting your current offer of X. If either you or your counterpart accept the other's active current offer of X, then that value of X determines both you and your counterpart's payoff in that round. If your offers agree or overlap, then the final value of X for that round is settled and your payoffs are determined by the midpoint of the two offers (e.g., if your offer is 460 and Player B's is 440 [your offer is 440 and Player A's is 460], for example, then the final value of X would be the midpoint of the overlapped values, or X=450 in this example).

You are not required to accept any of your counterpart's offers of X, nor is your counterpart required to accept any of yours. Should you reach the end of the round and not settled on a value of X (i.e., no current offer was accepted by either of you), then you will be prompted to input your final offer of what you would like X to be (and your counterpart will do the same). If your final offers agree or overlap, then the final value of X for that round is the midpoint. In the event no agreement is reached and your final offers do not overlap, both you and your counterpart will earn 0 tokens in that decision round.

Out of the 10 decisions rounds for this particular stage, we will randomly draw one round for payment at the end of the experiment (and each round is equally like to be the one that counts for payoff). At the end of the experiment tokens will be converted to cash. Remember, each token is worth \$0.10 or 10 tokens is worth \$1.00.

### **FOA Instructions (follow after the first 5 rounds)**

For the next 5 rounds, the following procedure will be used to deal with the possibility that you and your counterpart may not reach an agreement on the value of X by the end of the round (i.e., neither of you accepted the other's offer and your final offers did not come to agreement).

In the event no agreement is reached, the computer will choose between your final offer and your counterpart's final offer (recall, you will submit those at the end of the round and if they do not agree for a last chance settlement, then those final offers will be used in this settlement procedure). **In other words, the computerized settlement procedure will select either your final offer or your counterpart's final offer to be the settlement value of X for that round.** The computer may be more likely to choose some offers over others, but there is a random element to the computer's choice.

To give you some information about how the computer will select between you and your counterpart's final offer (in the event of dispute), imagine that the computer will first select its own value of X. The computer will then compare its value of X to you and your counterpart's final offers, and it will choose whichever of your final offers is closer to the value of X it (the computer) drew.

The table below shows 100 values drawn by the computer. From this, you can get a sense of which values of X are more or less likely to be drawn if this procedure is needed. So, if you and your counterpart cannot agree on a value of X, then the computer will draw a random value of X, it will compare the value it draws to you and your counterpart's final offers, and then the computer will choose whichever final offer (you or your counterpart's) that is closer to the random value it drew. As before the X-value for that round will dictate the payoffs that you and your counterpart would earn for that round.

|     |     |     |     |     |     |     |     |     |     |
|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|
| 477 | 593 | 482 | 405 | 552 | 419 | 561 | 439 | 503 | 491 |
| 417 | 519 | 467 | 460 | 499 | 498 | 404 | 535 | 391 | 538 |
| 544 | 571 | 442 | 488 | 438 | 506 | 550 | 473 | 539 | 439 |
| 515 | 529 | 458 | 589 | 403 | 518 | 519 | 564 | 641 | 534 |
| 435 | 593 | 571 | 550 | 537 | 447 | 420 | 406 | 486 | 494 |
| 525 | 645 | 394 | 516 | 617 | 470 | 483 | 430 | 392 | 515 |
| 421 | 545 | 449 | 456 | 562 | 464 | 460 | 460 | 599 | 382 |
| 453 | 515 | 418 | 452 | 537 | 446 | 503 | 558 | 515 | 483 |
| 480 | 553 | 412 | 575 | 451 | 509 | 594 | 382 | 464 | 511 |
| 442 | 562 | 463 | 467 | 524 | 596 | 483 | 424 | 528 | 520 |

At the beginning of each of the next 5 decision rounds, you will also be asked to indicate what you think the computer would likely select for its own value of X, should the computer draw be needed. If your answer is within 5 units of what the computer typically draws for X then you will earn an additional 20 tokens. Each new decision round repeats this process until new instructions say otherwise.

Remember, out of the 10 decisions rounds for this particular stage, we will randomly draw one round for payment at the end of the experiment (and each round is equally like to be the one that counts for payoff). At the end of the experiment tokens will be converted to cash. Remember, each token is worth \$0.10 or 10 tokens is worth \$1.00.

## APPENDIX C: Learning Trend versus Outcomes-based learning

### Motivation

Our approach to the main analysis uses a specification that considers dispute likelihood impacted by both a learning trend (captured by the *Treatment Round* variable) and the cumulative dispute history. An approach was developed in Bolton and Katok (1998) to more formally test whether using past outcomes captures the data better than a learning trend. Our approach in the main text has assumed that cumulative disputes as well as a learning trend may impact the data, but we did not address the hypothesis that all past outcomes (both settlements and disputes) motivate learning in the bargaining pair. Because our main text specifications include a *Round* variable to capture persistent learning and a *Dispute History*, or *DH*, variable, our approach is similar to capturing the impact of cumulative settlement history, or *SH*, via the *Round* variable. Specifically,  $Treatment\ Round-1 = SH + DH$ . In other words, the difference between *DH* and the *Round* estimate is co-linear with cumulative settlements history (for this reason, we do not include all three measures as co-variates in our regressions). Our Table 5 (main text) estimates of the narcotic effect are similar in spirit to Bolton and Katok (1998) in showing the additional explanatory power of the outcomes-based learning over a persistent learning trend.

A more direct test of outcomes-based over round-based learning that follows Bolton and Katok (1998) proceeds as follows: they define  $Outcome = DH - SH$ , which unlike  $DH + SH = Round-1$  will not be collinear with *Round-1*, and derive the following baseline estimation equation:

$$Dispute = \alpha + \beta(Round-1) + \gamma(Outcome) + e \quad (C1)$$

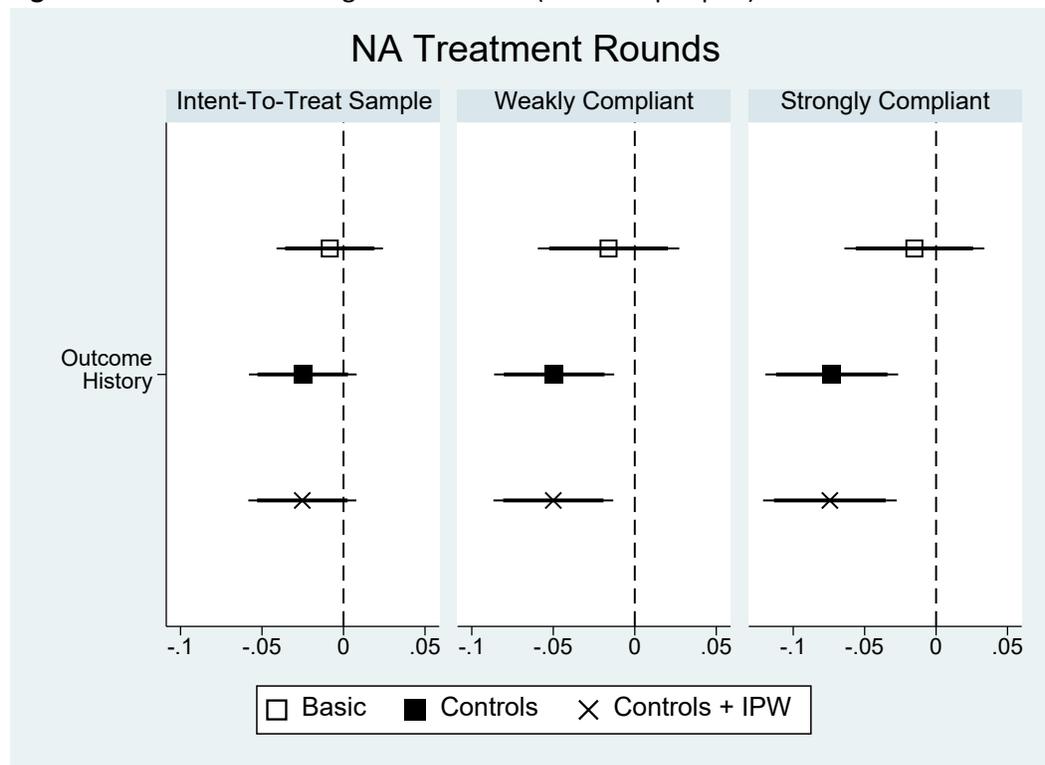
where *Dispute* is an indicator for dispute in the current round. From this, we test the hypothesis that  $\gamma = 0$ . Rejection of this hypothesis implies the outcomes-based model has better explanatory power than the persistent learning round-based model. We performed this test on our data and present the results below by means of coefficient plots of the key  $\gamma$  estimates, using model specifications as in the main text. The three panels show estimates from the different samples (intent-to-treat, weakly compliant, strongly compliant bargaining pairs), while the different rows of estimates reflect the most sparse model specification (*Round-1*, *Outcome*, *#SR*, and *BeliefGap* in the case of the FOA treatment), a model with more full control variables (as in main text analysis), and a model that corrects for sample selection using the pair-average inverse probability weights (as in main text analysis). Figure C1 shows this for the subset of only the NA treatment rounds, Figure C2 shows the results for the FOA treatment rounds, and Figure C3 pools all rounds together.

As can be seen in these coefficient plots, the outcome history is not uniformly statistically significant, but it does appear to be so for all the models with additional co-variates that condition on some level of compliance by the bargainers. The statistically significant and *negative* coefficient estimates on *Outcome*, which measures cumulative net-Disputed rounds, indicates that more history of dispute relative to settlement tends to decrease the likelihood of dispute in the current NA bargaining round. In contrast, in Figure C2 it is clear that we uniformly reject the round-learning in favor of the

outcomes-based learning in all specifications and for all samples. And, consistent with our findings reported in the main text, a higher level of net-dispute in a pair’s bargaining history in FOA increases the likelihood of current round dispute (note: this is the main narcotic effect as opposed to the SR-induced additional narcotic effect. Here our goal was to test the outcome-based learning model in the general sense). The same is found in Figure C3 that pools the two treatments together.

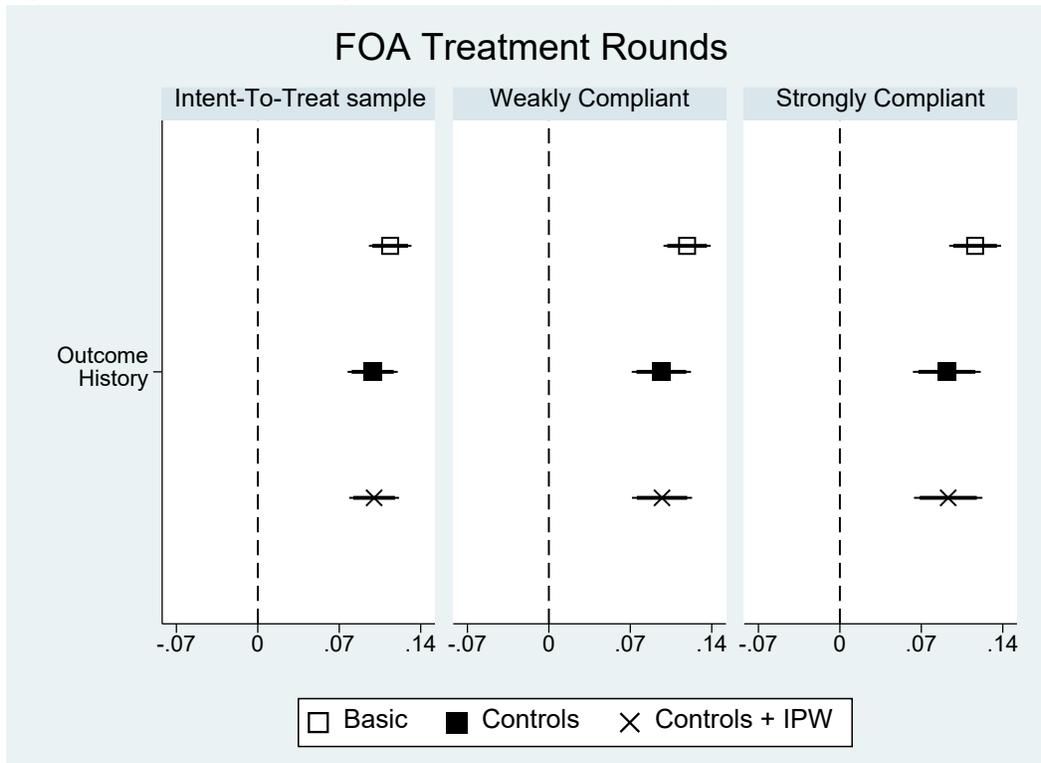
In short, based on this more formal test, an outcomes-based learning model has more explanatory power in our data, as was the case in Bolton and Katok (1998). Additional estimations (not reported here, but available on request) also find that when we include both cumulative dispute and cumulative settlement history in the specification, current round dispute is more effected by history of dispute (so our main text analysis focused on this). However, our results support the notion that increased history of dispute actually *increases* the likelihood of dispute. This is in contrast to the Bolton and Katok (1998) result found in their somewhat different arbitration environment where they concluded past disputes slowed the rate of learning with arbitration. Our result is consistent with results in Dickinson (2005) that found a dependency-effect regarding a more comparable FOA laboratory procedure.

**Figure C1:** Outcome learning estimates: NA (5 rounds per pair)



**Notes:** coefficient plot shows point-estimates of the variable Outcome in the specifications based on equation (C1) above. The thin (thick) lines represent the 99% (95%) confidence interval on the estimate assuming a 2-tailed test. Positive (negative) estimates on outcome history indicates that more disputes relative to settlements in the pair’s history tends to increase (decrease) the likelihood of disputing in the current round.

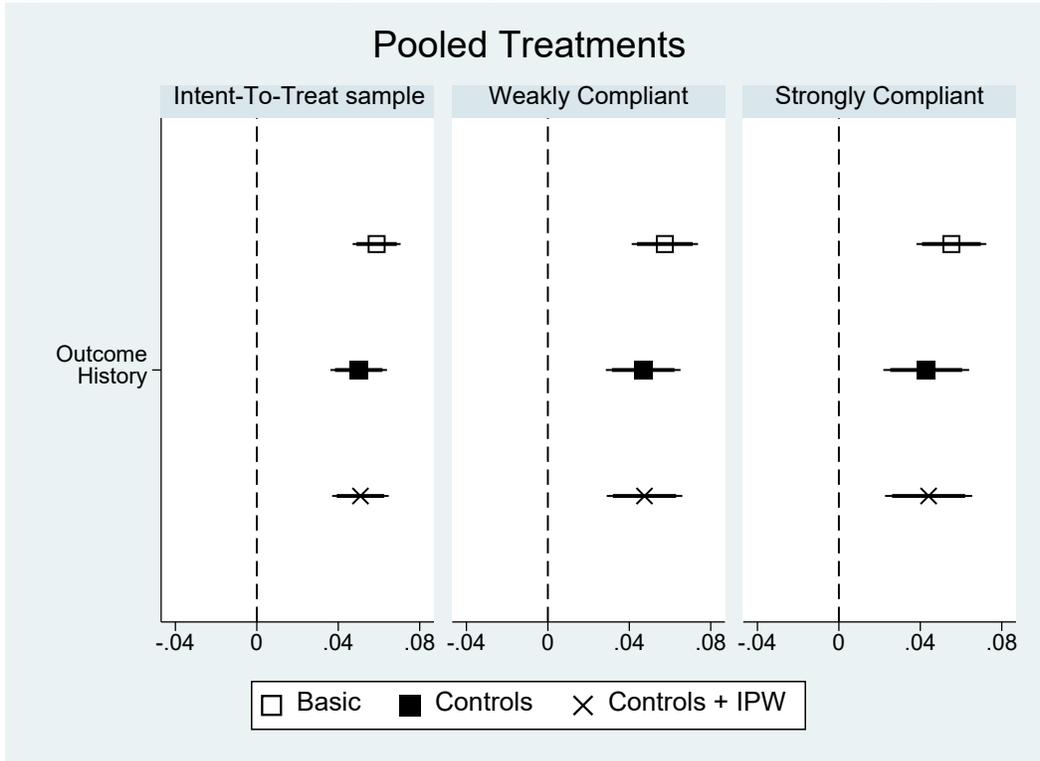
**Figure C2:** Outcome learning estimates: FOA (5 rounds per pair)



**Notes:** coefficient plot shows point-estimates of the variable Outcome in the specifications based on equation (C1) above. The thin (thick) lines represent the 99% (95%) confidence interval on the estimate assuming a 2-tailed test. Positive (negative) estimates on outcome history indicates that more disputes relative to settlements in the pair's history tends to increase (decrease) the likelihood of disputing in the current round.



**Figure C3:** Outcome learning estimates: pooled treatments (all 10 rounds per pair)



**Notes:** coefficient plot shows point-estimates of the variable Outcome in the specifications based on equation (C1) above. The thin (thick) lines represent the 99% (95%) confidence interval on the estimate assuming a 2-tailed test. Positive (negative) estimates on outcome history indicates that more disputes relative to settlements in the pair's history tends to increase (decrease) the likelihood of disputing in the current round.