

DISCUSSION PAPER SERIES

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Ian Gazeley

University of Sussex

Rose Holmes

University of Sussex

Andrew Newell

University of Sussex and IZA

Kevin Reynolds

University of Sussex

Hector Gutierrez Rufrancos

University of Stirling.

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ABSTRACT

Inequality among European Working Households, 1890-1960*

In this article we map, for the first time, the time-path of the size distribution of income among working class households in Western Europe, 1890-1960. To do this we exploit data extracted from a large number of newly digitised household expenditure surveys. Many are not representative of the population, or even of their target-subpopulation, as methods of social investigation were initially primitive, though rapidly evolving over this period. We overcome the consequent problem of comparability by exploiting our knowledge of the methods used by early social investigators to estimate of the scale of known biases. For some we have the original household data, but in most cases we have tables by income group. One by-product of this work is an evaluation of the range of estimation methods for distributional statistics from these historical tables of grouped data. Our central finding is that inequality among working households does not follow the general downward trend in inequality for the early part of the century found in labour share and top income studies. Contrary to Kuznets' prediction, our evidence suggests that on average income inequality among European working households remained stable for three generations from the late nineteenth century onwards.

JEL Classification: N33, N34, O15

Keywords: inequality, working households, Europe, 20th century

Corresponding author:

Andrew Newell
Department of Economics
University of Sussex
Brighton, BN1 9SL
United Kingdom
E-mail: a.t.newell@sussex.ac.uk

* Funded by ESRC grant ES/L002523/1 'Global Income Inequality, 1880-1960', February 2014-January 2018.

Introduction

This article is an attempt to estimate the paths of the size distribution of income among working households across Western Europe, 1890-1960. The history of income inequality might seem to be settled, as there exists a wealth of previous work on this topic, see for instance Morrisson (2000), Williamson (1985, 1999), and the work of the top incomes group presented at The World Wealth and Income Database (WID.world). Indeed, Lindert (2000:12) argued persuasively that it might be advisable to move away from trying to establish trends:

...simply asking whether or not a country follows the famous inverted-U slows us down, by delaying our search for the more interesting interplay of underlying forces that give us a rich history of episodic movements, not just a long up-and-down movement. Given the opportunity to explore the changing effects of government policies, laws, wars, demography, technology and other forces on inequality movements from one episode to another, why settle for a debate over a single curve?

So, why do we seek further evidence? It is the first, necessary step for us to follow Lindert's suggestion of moving on to the analysis of change. We bring new data, having collected and digitised data and statistics from around one hundred and forty northern European household expenditure surveys. Household expenditure surveys form the major data resource for the contemporary study of inequality, but very little of the existing historical literature is based on such data. Instead, most historical studies rely on tax data, as well as piecing together evidence from key historical time series such as average wages by skill and labour's share in national income.

Inferences about the path of inequality from tax data, as Morrisson (2000) make abundantly clear, suffer from the vagaries of tax reform, especially in a period that saw the rolling out of more-or-less universal income taxation. Similarly, inferences about inequality made in studies based on wage or sectoral relativities tend to rely on assumptions on the size and direction of movement of within-skill-group inequality. Neither tax data nor pay relativities can help unpack how changes in income sources influence, and are influenced by, household formation and the allocation of paid work within the household. This is especially important for the study of inequality in Europe during the first half of the twentieth century, because average household size was smaller by mid century than it had been in the late nineteenth century, though the magnitude of this decline could vary by country, region, socio-economic

class, religious affiliation and ethnicity. By contrast, individual household survey data offer the chance to observe these changes and their impacts on inequality, which is why they are the preferred data source for modern measures of inequality.

Of course, inference on inequality from historical household expenditure surveys is not without problems. One well-known reservation about using household expenditure data sets is that many are not representative of the population, or even of their target-subpopulation, as methods of social investigation were initially primitive, though rapidly evolving over this period. We tackle the issue in two ways: first, our archival work expands considerably upon the volume of digitised datasets, which inevitably helps to alleviate the representativeness problem. We also confront the problem head-on by exploiting our knowledge of the methods used by early social investigators to estimate of the direction and scale of methodological biases.

The analysis of tax data has established (see WID.world) a clear picture of declining top incomes in many countries across the half of the 20th century, but trends in lower parts of the income distribution over this period are not yet established. There is plenty of evidence of improving living standards in Western countries, but we know little detail about how the relative incomes of poorer households fared and what the key drivers of change were. There is a long list of potential influences, including: technological change that may have shifted the demand for labour across the skill spectrum and changed the nature of unpaid domestic work; improvements in education; improvements in the social safety net; the demographic transition that resulted in fewer mouths to feed per household and allowed adult women to choose market work for more years in their lifetime; and greater trade union organisation of less-skilled workers.

At present, there is very limited statistical coverage of household inequality prior to the Second World War. For national inequality measures, there are four main international collections. The largest is the World Income Inequality Database (WIID) for the World Institute for Development Economics Research (WIDER). This gives, at the time of writing, over 8,800 Gini coefficients, as well as other inequality indicators, from around the world, and dating back as far as 1867. But only 10 Gini coefficients in the set are for West European countries prior to 1940 and only 24 prior to 1950. After that, coverage picks up but still there are only 59 for West European countries prior to 1960. Therefore, this database is not extensive enough to analyse changes in distribution within West European countries

before 1950. None of the other collections, from the OECD, the Luxembourg Income Study and the World Bank's Povcal.Net, give measures earlier than 1980. An exception is the OECD study by Sawyer (1976), which offers Gini coefficients for some West European countries for the late 1960s and early 1970s.

Our aim here is to contribute to the filling of these gaps in our understanding of an important period in world economic development: rapid technological change was driving sustained increases in per capita incomes in Western Europe, but progress was also riven by two global conflict and an intense depression. The structure of the paper is as follows. In the next section, we briefly place our research in the context of existing estimates on inequality for Western Europe, though this is not intended to be an exhaustive literature review and concentrates on evidence relating to the bottom end of the distribution in our period. Section 3 contains a discussion of the evolution of household expenditure surveys over this period, in respect of their methods and purpose. In section 4, we evaluate, with illustrations, the various ways the distribution of income/expenditure between households and people can be measured. After that, we discuss our methods of estimating inequality from grouped data, referring the reader to a working paper for detail, and offer some illustrative results. Finally, we study how our estimates vary systematically, depending on the nature of the underlying data, on the method of estimation, and other features of the data. This allows us to answer our research question: is there a significant Pan-European trend in inequality among working households over the early part of the twentieth century? Perhaps, surprisingly, we do not find one.

2. Our approach in context

The starting point for the literature on trends in economic inequality is Kuznets' (1955) seminal work that posited an inverted U-shape pattern in the time path of inequality. For Kuznets, inequality rose during early phases of industrialisation before falling as the economy matured. Kuznets' work influenced a generation of scholars who sought to investigate the empirical basis for his hypothesis. Where household income/expenditure data are unavailable, researchers have developed ways of estimating inequality from data that are available. Since the Theil index of income inequality is decomposable into income inequality between and within incomes from different sources, one approach is to add up these

components (see Morrisson, *op. cit.*, p219). Key indices in this approach are: the share of labour in national income, earnings/wage differentials by skill, and tax revenue data.

Acemoglu and Robinson (2002), summarising the key cross-country studies of Williamson (1985) and Morrisson (2000), state that income inequality, as measured by the Gini, rose in the latter part of the 19th century and then fell through the early decades of the 20th century in the UK, France, Germany and Sweden. However, in the Netherlands and Norway, inequality fell continuously from the mid-nineteenth century onwards. Williamson (1985) argues that the available historical data for the nineteenth and first three-quarters of the twentieth century in the UK, USA, Sweden, Denmark, Netherlands, West Germany and Prussia, conforms to the downside of the Kuznets curve (see also Williamson 1991:58-59). Overall, the evidence for the rise in inequality during the initial phase of industrialisation - which is fairly robust for the UK and USA (see Lindert, 2000) and more tentative for nineteenth century France (Morrisson and Snyder, 2000) - is less widespread than the evidence for declining inequality during later phases of economic development, until around 1980, at least.

Most studies that have assessed changes in the functional distribution of income over time, find the share of wages in national income increasing fairly continuously over the first part of the 20th century (Kaelble and Thomas, 1991:35). However, most of these economy-wide estimates begin in the interwar period or later. For the nineteenth and early twentieth centuries Phelps, Brown and Browne (1968) calculated wage/income ratios 1860-1960 for the UK, USA, France, Germany and Sweden, defined as the average industrial wage relative to industrial income per occupied person, as a means of approximating 'the division of net pay between pay and profit' in the industrial sector of the economy (1968: 60-61). Dividing the century from 1860 to 1960 into three periods (before 1913, interwar and post 1950), they found: increasing wage/income ratios in each period for the UK; increasing during the interwar period and then remaining roughly stable for the USA; increasing only in the post-war period for Germany; and fairly flat across all three periods for Sweden (1968: 365 Table 32). They were not able to estimate wage/income ratios for all periods for France, though for the four other countries the wage to income ratio remained around 0.6-0.7 until after the Second World War (1968: 336).

Analysis of income inequality in the lower half of the distribution has also concentrated on the behaviour of pay ratios of male adult unskilled to skilled workers. According to Williamson (1991:62), economy-wide pay ratios for the UK increased during the first half of

the nineteenth century, reaching a peak in 1851, before falling back continuously thereafter to 1911. Horrell and Humphries (1992:863) using household expenditure survey evidence confirm this rise in the early 19th century 1787-1865. They find that the increase in adult male earnings inequality by skill is also reflected in a rise in household income inequality. From about 1914 there is a well-documented decline skilled wage premia among manual workers, see, for instance Gazeley *et. al.* (2017).

For Germany, Dumke (1991:131-2) finds falling pay ratios during the late nineteenth century to the turn of the century, rising again to 1914, then falling through WW1 and the Weimar period, before a short-term rise through the Nazi era, followed by substantial falls in the 1950s (1991:128 Fig. 5.1 (b)). In Sweden, Solderberg (1991:82-3) argues that adult male skilled/unskilled daily pay ratios for Stockholm follow a similar pattern as the one identified by Williamson for the UK, with a peak around 1840-5. However, economy-wide data that take into account composition effects show rising inequality 1870-1914, followed by strong levelling across the First World War until 1920. Inequality measured in this way, then widens again during the deflation of the 1920s and across the Great Depression, before a sustained and dramatic levelling from the early 1930s onwards, which continues through the Second World War to 1950 and after.

Narrowing nineteenth century wage differentials are also found in Belgium by Scholliers (1991) between the wage censuses of 1846 and 1896. However, it is noteworthy that Scholliers is also able to compare estimates of adult male pay inequality with near contemporaneous estimates of inequality from Belgian household expenditure surveys. He shows that inequality among households was considerably less than inequality among adult male wage earners at both dates, but especially so in 1896. He notes that at mid-century there were:

...different kinds of subsistence strategies in families where the breadwinner earned a low wage. Female and child labour, small scale shopkeeping, taking in lodgers and so on were all means by which the family income was increased. This 'topping up' income led to a slight decrease in inequality in terms of family income. This is illustrated by the figures for Ghent working class family incomes in 1853..... Total family incomes were 75 per cent higher than those of the head of the family in a low wage category and 52 per cent higher in the high wage category. The pay differential between the respective breadwinners amounted to nearly 100 per cent, whereas family income differentials amounted to 'only' 70 per cent (1991: 103-4).

By 1896, the practice of ‘topping up’ low wage head of households’ income from other sources had a greater effect on measured household inequality, reducing differences in household incomes by skill of principal wage earner to ‘virtual insignificance’ (1991:111-2). This is an important result in the context of our research, as it suggests that the evidence of generally falling West European pay ratios between the late nineteenth century and mid-twentieth century, may not be as evident in household-based measures of inequality.

The most influential set of studies of the last fifteen years or so exploit tax revenue data to estimate the income shares of top income groups. It is well known that regular survey methods are not good at collecting information from very wealthy individuals, so the addition of long-run trends in these shares offers much new information. This has been the work of a group, led by, among others, Facundo Alvaredo, Anthony Atkinson, Thomas Piketty and Emmanuel Saez. The results, and reference for the research papers on which they are based, are collected together in The World Wealth and Income Database (WID.world). This database contains trends in top income shares for many European countries. In brief, for the first half of the 20th century they report declines in the top income shares in Denmark, Finland, Norway and Sweden. Similarly they report declines in top income shares over the same period in France, Germany (to 1939) the Netherlands and the United Kingdom. Tax data are not yet available for Italy, Spain and Portugal, so knowledge of South Western Europe remain mostly out of reach. Below we discuss the relationships between our findings and the results of the top incomes work.

Lastly, Rossi *et. al* (2001) is an ambitious attempt to build a household survey-based statistical picture of the path of inequality in Italy 1881-1961. Their research takes a very different approach to that adopted here, which is described below in detail. Briefly, our approach is to estimate household-level inequality within the existing surveys. The surveys, with a range of different objectives, ask different sets of questions, so that, for instance, we know little about the people in the household in some surveys, or we only have income data, or expenditure data in other surveys. Surveys where there is insufficient information to estimate any form of inequality are not considered, and cases within surveys with insufficient information are dropped. We estimate measures of inequality, made on a variety of bases, and then, later, gather these into a data set that also contains indicator variables for the various aspects of the surveys (time, place, target population, the use of income vs.

expenditure etc.). This allows us to control for the impacts of differences in the surveys on measures of inequality.

By contrast, Rossi *et. al.* (2001) set themselves the task of creating a data set that is as close to what is known of a population as possible. They use existing budgets to create a target survey structure, to which only a subset of existing surveys or budget observations, can be made to conform. In cases where information is incomplete, they employ a hot-decking procedure to use observations (households) where data exist to ‘donate’ information to observations where data do not exist, thereby creating a set of synthetic observations (2001, p. 909). Subsequently, they make adjustments for population representativeness. In summary, the approach of Rossi *et. al.* is to try to augment the existing fragments of data and, finally, to create a synthetic population. The approach taken in this paper, by contrast, eschews the creation of synthetic data and relies on regularities between results from sources of similar types to allow inference about longer-term movements.

3. The Evolution of Household Budget Surveys in Western Europe, 1850s-1950s

Household budget surveys seem to have been a West European innovation, and their evolution is implicitly involved with the emergence of particular conceptions of West European nation-statehood and citizenship. We see, between the 1850s and 1950s, a four-stage shift in the methodology and focus of domestic household budget surveys, though each stage is not necessarily temporally unique or self-contained. The Western European nations involved in colonial expansion developed separate techniques for analyzing the domestic structure of the family lives of citizens of overseas territories, which we do not consider here (see Gazeley and Holmes 2018 for a fuller description).

The first modern attempts to collect and analyse household budgets that we are aware of are due to David Davies (*The Case of Labourers in Husbandry*, 1795) and Fredrick Eden (*The State of the Poor* 1797). Eden and Davies collected household budgets of income and expenditure from the rural labouring poor in the late eighteenth century for the explicit purpose of providing empirical data to inform contemporary debate on the cost of poor relief and reform of the poor laws (Gazeley and Verdon, 2014). According to Stigler (1954) two developments led to a revival of interest in budget studies from the mid-nineteenth century onwards: the wave of unrest that swept Europe culminating in the revolutions of 1848 and

developments in statistical theory that permitted a more sophisticated analysis of social data (1954:96). Household budget surveys in the latter half of the nineteenth century were mostly carried out from a sociological perspective, either by social reformers or academics tending to follow the LePlay/Ducpétiaux model of detailed social analysis of households alongside a breakdown of income, expenditure, food consumption and nutrition.¹ International collaborative endeavour between researchers began to shape a shared methodology for conducting surveys in a series of nine meetings of the International Statistical Congress (ISC) held between 1853 and 1879 (Randeraand, 2011). The Belgian statistician Adolphe Quételet and other European statisticians were, during this period, increasingly turning their attention to the use of statistics to measure human growth and development (Eknoyan, 2007). Household budget surveys also provided data that allowed the formulation of statistical regularities such as Engel's Law. While the ISC meetings helped formulate international statistical knowledge-sharing, Randeraand (2011) argues that their main function was to contribute towards the acceptance of statistics as a fundamental guide in the political decision-making of nation-states.

The second stage of development focussed on collecting comparative international data, primarily to evaluate the impact of tariffs and trade. In the United States, the Commissioner of Labor, Carroll D. Wright, conducted an extensive survey of household income and expenditure in the US and five West European countries in 1889/90, motivated by the McKinley Tariff question (Williamson, 1967). Before the turn of the twentieth century, West European governments also began to see household budget research as part of international trade and labour research. From the 1890s, departments of international labour began to undertake research employing methodologies adapted from those used by social investigators and academics. In the first decades of the twentieth century, most European governments created statistical offices, and the responsibility for undertaking domestic household budget surveys shifted to state-employed statisticians. Perhaps the best example of this type of enquiry is the British Board of Trade's collection of household budgets (1908-1911), which investigated the cost of living in France, Germany, Belgium and the United States, for the purpose of comparison with the UK. The Board was influenced by Le Play's mid-nineteenth

¹ Key examples include: F. Le Play, Alfred Mame Tours et fils, *Ouvriers européens. Études sur les travaux, la vie domestique et la condition morale des populations ouvrières de l'Europe, précédée d'un exposé de la méthode d'observations par*, 2^e éd., 1877-1879. 6 vol.; E. Ducpétiaux, *Budgets économiques des classes ouvrières en Belgique* (1855); B.S. Rowntree, *Land and Labour: Lessons from Belgium* (London: MacMillan and Co., 1910); Ernst Engel, *Die Lebenskosten Belgischer Arbeiter-Familien Früher und Jetzt* (Dresden: C. Heinrich, 1895 [1857]).

century investigation of working-class living conditions in Europe, and the later investigations of Belgian household budgets by Ducpétiaux as well as the German budgets conducted by Gruber. From their inception, government-instigated household expenditure enquiries in the UK were motivated by two objectives: the desire to calculate a cost of living index and the comparative analysis of living conditions among industrial competitors.

We see a third stage beginning after the First World War, as European and non-European governments, guided by the League of Nations, became increasingly invested in developing shared methodologies of budgetary surveys – primarily for the purpose of constructing cost of living indices. Held under the auspices of the International Labour Organisation (ILO), the 1926 International Conference of Labour Statisticians was attended by representatives from almost every European nation and several non-European nations. They agreed that household budget surveys were the best way of assessing living standards, and adopted a clear, shared methodological framework that if possible, countries would use in their investigations (International Labour Organisation, 1926). It also, for the first time, agreed on a timescale with countries that had not undertaken an enquiry since 1920/21, resolving to launch major surveys using the new methodological approach before 1928. In the same (interwar) period, however, it is worth noting that some prominent researchers, notably Carle Zimmerman (1936: 378-415), felt that something had been lost in the move towards internationalism, and that regular, local, small-scale representative studies taken very much in context were the best way to ensure accurate reporting.

The final stage is best seen as driven by a desire to implement developments in statistical methodology to improve the representativeness of samples in relation to their target populations. In 1949, the ILO again prioritised household budget surveys and distributed new methodological guidelines to member states, emphasising the importance of random sampling, accurate record-keeping and international uniformity (International Labour Organisation 1949). Having previously focused chiefly on urban, working-class nuclear family households, survey-takers now employed random sampling techniques to attempt representative population analysis. In the late 1940s and 1950s, surveys thus became more comprehensive, partly due to the continued international co-operation between the statistical branches of European governments, and partly due to the new interventions of the UN, which emphasised household budget research on a global scale (Darmois *et al*, 1949).

Household budget surveys emerged as a way for social reformers to demonstrate the condition of the poor and, by implication, to influence contemporary debate on social policy. In the nineteenth century, they were the petri dish of statistical methodology. By the turn of the twentieth century, they were an accepted apparatus in both the domestic economy and foreign policy decision-making process of nation-states. And by the mid-twentieth century they allowed NGOs to hold nation-states to account for the comparative household economies of their citizens. Always products of their time, they demonstrate the emergent power of crisis-born institutions and the move away from nineteenth century forms of nationalism towards a globalised economy. Prais and Houthakker (1971) argue that the value of household budget studies is to comment on the condition of the people, derive weights for cost of living indices and for empirical studies of consumption behaviour (1971:4). We would add to this list their value to the investigation of inequality among the working class families surveyed.

This study exploits the set of European household expenditure surveys collected by the Global Income Inequality Project². Although we have recovered and analysed individual level data from around forty household surveys in this period, the majority of the items in the collection are tables from survey reports, rather than individual-level data. These tables present averages from the surveys grouped by household income. They contain the number of households in each group, as well as, variously and among other things, averages of: income, expenditures on main items and numbers of children. Table 1 presents the coverage of our findings, by country and decade. The total number of surveys found by the project is much larger, but here we count only those that provide useable data or statistics. There are, for instance, eighty-one surveys that are usable for the estimation of inequality for the pre-1940 period. This is a major advance on WIDER's WIID database.

<Insert Table 1 here>

4. Estimating inequality from household expenditure surveys

There are three main challenges in this regard. First there is the choice of unit of measurement: income, expenditure, per household, *per capita* etc. Where we have individual household-level data, the surveys vary in the range of questions asked of respondents but

² Funded by ESRC grant Global Income Inequality 1880-1960, grant no. ES/L002523/1

usually several possibilities exist. Where we have grouped tables, the unit of measurement of the group is what we have to use, most commonly, total household income. Thus, limitations to the available information on household composition affects our choice of the key unit of measurement. Second, there is a set of issues surrounding the creation of parameter estimates from income-grouped tables. Third, for estimates derived from either grouped tables or individual-level data, there is the question of population representativeness of the samples, and the possible adjustments that can be made. Here we take a regression approach by treating the sampling characteristics of a survey as explanatory variables in regression seeking to explain differences between survey inequality measures. This is further discussed in Section 6 below. Here we discuss the first two challenges.

4.1. The choice of unit of measurement

There are a number of ways of calculating economic welfare among members of a household. We divide our discussion into three parts: the choice of numerator, the choice of denominator and whether we estimate inequality between households unweighted or weighted by family size. For the numerator, our surveys will typically offer us at least one of the following over the relevant period of observation: total household food expenditure, total household expenditure, or total household income. We find quite similar results where we can compare inequality estimates using either total expenditure or total income. Initially, this was surprising given that saving behaviour and commonly observed inconsistencies in responses to income and expenditure questions were likely to cause differences. This might come about because of the imperfections of recall while attempting to record expenditures and incomes as part of completing the questionnaires. Here we present, as a first step, estimates derived from all three measures.

The choice of the denominator, however, has a profound impact on our results. The contemporary standard method is to divide income or expenditure by the number of equivalised people, using an equivalence scale, such as the OECD Modified Scale.³ This requires knowledge of the numbers of people of different ages or, at a minimum, the numbers of adults and of children in the household. Where we have individual household data this information is often available. Where we have grouped data, this is problematic (see section 3 below). With individual household-level data we can employ Stephen Jenkins' (1999)

³ <http://www.oecd.org/eco/growth/OECD-Note-EquivalenceScales.pdf>

Stata suite of distributional statistics, for instance *sumdist* and *ineqdeco*. These give a wide range of inequality measures and have been used in scores of studies.

Lastly there is the issue of weighting by the number of people in the household. This is not always possible, but in cases where it can be done it makes sense to weight estimates of *per capita* income by the number of people, or estimates of incomes per equivalent adult by the number of equivalent adults.

4.2 Estimating inequality from grouped data

In most cases where we rely on grouped tables, households are grouped by total household income. There are cases where we have tables giving averages for households that are grouped by *per capita* or equivalised incomes or expenditures. For instance, many early Belgian and Norwegian studies present results grouped by equivalised incomes. For such cases there are various existing ways of estimating measures of inequality. Perhaps the simplest and certainly the best-known method is to employ Chen, Datt and Ravallion's Povcal⁴ software which is designed to estimate poverty and inequality from grouped statistics via the estimation of General Quadratic and Beta Lorenz curves, originally formulated by Villasenor and Arnold (1984, 1989) and Kakwani (1980), respectively.

This approach, however, has its downside for our study, since the Gini coefficient is Povcal's sole distributional statistic and this is very restrictive in terms of inequality measurement. So instead, we programmed our own suite of methods. Gutierrez Rufrancos and Newell (2017) describes and tests these methods, which are also summarised in Appendix 2. Here we have chosen to illustrate our findings using two alternative methods of estimation. Firstly, the 'Groups Naïve Frequency Method', simply assigns the group mid-point to each household in the group and then calculates the Gini coefficient. Secondly, we estimate the parameters of Kakwani's Beta-Lorenz approach (Kakwani, 1980). It becomes clear through the tables of descriptives and regression analyses that the two methods give very similar results.

5. Inequality measures

4

□ <http://go.worldbank.org/YMRH2NT5V0>

We constructed around 150 sets of inequality measures from our surveys (some surveys contained more than one sub-survey). Of these we obtain just over 130 that come with sufficient ancillary information for our estimation. Tables 2 and 3 give summaries of inequality measures estimated, in turn, by our naïve and Beta-Lorenz methods. The first thing to note is that since these data sets mostly contain information on working class households, measured inequality is generally much lower than it would be across the population. Secondly, note that the two tables give very similar information, except that while inequality as measured by Beta-Lorenz is around 10% higher than the naïve estimates, 90/10 and other decile ratios are lower when estimated by Beta-Lorenz. This comes about as the curvature in Beta-Lorenz assumes less density in the tails of the distributions than simple interpolation, but also swings lower though the centre of the distribution. Other facts that emerge are that *per capita* income tends to be less unequally distributed than household income. This is because, for many cases, household membership and household income co-vary positively.

<Insert Table 2 and 3>

Statistics derived from tables grouped by income tend to generate higher measured inequality, whereas tables grouped by characteristics other than income, such as occupation or household size, understandably, give much lower estimates. Post 1939 inequality is higher in these tables but this may be related some aspects of surveys that change over time, in particular, as discussed in section 3, sampling methods. Taking the surveys as a whole we see that those that followed stratified random sampling give much higher average levels of inequality. Of course there are other reasons, not tabled here, that would give rise to different levels of inequality. Firstly, sample sizes vary from scores to tens of thousands. Secondly, the target population varies across these surveys. For example: some are exclusively urban; others concentrate on manual workers' households only.

Table 4 studies one key conflation, for our purposes, by presenting mean Gini coefficients by era and by sampling method. Separating underlying trends in inequality from trends in sampling is the main objective of this study. The result is that post-1939 Ginis are only higher if they are known to derive from stratified random sampling. Other post-1939 Gini coefficients are not notable higher. This suggests the upward overall time trend in inequality in these statistics is likely to be a product of evolving sampling method. For the case of the UK, we have studied this issue in detail and confirm that there is no rise in inequality among

working households in this period once changes sampling methods and target population are taken into account, see Gazeley *et. al.* (2017).

6. Analysing the historical path of inequality among working households

In this section, we take estimated inequality measures and analyse the extent to which their variation over time and space is due to differences between countries, changes over time and differences in: sampling methods, target samples, and choice of groups. There is another potential source of variation: in Appendix 2 we describe how we tested six methods of estimating inequality from grouped tables. Early experiments showed us that the important results with respect to non-estimation issues were unaffected by the choice of estimation method. This results is foreshadowed by the similarities in relative inequality measures in Tables 2 and 3 where results were presented for, respectively, our ‘naïve’ and Beta-Lorenz estimation methods. Here again, to illustrate this without asking the reader to wade through too many tables, we present our key findings from regressions in which countries are pooled and the dependent variables are statistics estimated two different ways: by our ‘naïve’ and Beta-Lorenz methods, see Appendix 2 for a longer discussion.

In the pooled regressions given in Tables 5 and 6, the Gini coefficient, the 90/10 and 50/10 ratios are the dependent variables. The list of explanatory variables includes (unreported) country controls, and controls for the era in which the surveys were carried out. We choose break point years 1918 and 1939, though results are not sensitive to the exact choices of years to separate the interwar years from before and after.

In the first column of Table 6 the dependent variable is the Gini coefficient calculated, as discussed above, by our simple naïve method. We investigate how this varies with all the measured and recorded characteristics of the surveys by weighted least squares, with the weights being the square roots of the numbers of respondents in the surveys⁵. Taking the explanatory variables in turn, we find that including rural households in the survey generally

⁵ This is a standard correction for heteroscedasticity in cases like ours, see a very useful recent discussion in Solon *et. al.*, (2015).

raises the estimated Gini, as does the inclusion of urban households. The two coefficients are close in size, giving us the result that a survey that covers both rural and urban areas, rather than only one of these, will provide an estimated Gini, on average, between 8 and 9 points higher.

Next we control for cases where the survey tables are not simply delimited by income. From an income perspective these cases might create tables of overlapping income groups, and so one would expect a lowering of measured inequality if we treat them as non-overlapping. To be clear, if households were randomly assigned to groups, then there would be only random differences in group average income, so that estimated between group inequality would be negligibly small. But most of the groupings: region, social class, household size occupation, do indeed vary quite substantially with income. Here we find a strong negative effect reducing inequality, approximately eleven Gini points, if non-income groups are used. Another case is where we are given income groups, but by one or more separate characteristics additionally. In other words there is more than one group in the same income range. This should not have the effect of lowering Gini inequality, and we do not find a strong negative or positive effect on the Gini.

Next we control for cases where the survey tables have not given us the within-group mean income. Here we simply replace the mean with the group mid-point. This replacement could have different effects in different parts of the distribution, but the largest effects will be at the extremes, where the mid-point of a group would be further from the centre of the distribution than the mean and thus using it would bias a Gini upwards. The results support this with a substantial average rise of just under five Gini points. We also control for cases where only households headed by a manual worker are considered, but this gives us no notable results.

Another set of controls reflect the survey type: stratified random sample, snowball aiming at a particular group; snowball attempting to be representative; or unknown. The unknown are the omitted group. Perhaps surprisingly, none of these seem to generate significant average differences. These results should be taken together with results from the era variables, since we know that sampling methods were evolving. We turn to these next.

As Table 4 shows, if one compares mean Ginis by era *and* method of sampling the only group with a much higher mean is the group of Ginis from post-1939 stratified random samples. The other surveys in the period do not have a higher average Gini, and there were

very few surveys in our collection, just three, conducted by stratified random sampling prior to 1939. There were, however a larger number of non-stratified random samples among the post 1939 surveys. We test for this by including an interaction between the indicator for a post-1939 survey and one for a stratified random sample. Though this is not significant, the negative coefficients on the pre-1939 era variables are negligible and conventionally insignificant. Thus this regression confirms that the Table 4 result remains in place after many controls are added. The other columns for Table 5 show very similar results for the 90/10, 90/50 and 50/10 decile ratios. Across 40 different measure of inequality, we find no evidence of a region-wide trend in inequality among working households. Of course one might argue that the estimation method for the inequality measure might be so noisy around ‘true’ inequality that not much can be inferred, but we do find a number of sensible results in this table, on the geographical extent of the surveys, types of groups used and so on.

The point of the results in Table 6 is to repeat the analysis of Table 5 with Beta-Lorenz estimated inequality measures. Here the key results are closely in line with those of Table 5. Again we find no sign of a general downward trend in inequality. However here we find that interaction term that singles out post-1939 survey with a stratified random sampling design, is well determined and negative in two cases out of four (the 90/10 and 50/10 percentile ratios

<Insert Tables 5 and 6 here>

7. Conclusions: Inequality among working households 1890-1960

Any attempt to explain estimated variation in income inequality measures over time and space need to be able to identify the extent to which these are due to genuine differences between countries, or due to differences in sampling methods, target samples, and choice of groups. In this article we go some way towards resolving this conundrum. We have built up a database of household expenditure surveys for Western Europe 1890-1960 that extends the existing evidence base by about a factor of ten and charts the behaviour of inequality, for the bottom 4/5ths of the population (more at the end of the period). The results of the analysis of over 130 surveys warn very clearly of the pitfalls of inferring changes and other comparisons based upon inequality measures created with differing methodologies. On the other hand,

there are regular-looking patterns in measured inequality, for instance, the effects of: limiting regional coverage; differences in the grouping of cases and from the use of midpoints in place of group means.

The most notable outcome of this analysis is the lack of a definite long-term trend in these inequality measures. This is a surprising result because other indicators of inequality from the lower half of the distribution suggest a clear trend of declining inequality during the twentieth century for most West European countries. Our household-level results suggest that household inequality may not have followed the generally declining path of inequality on the downside of the Kuznets curve that tax, national accounts and relative wage data show.

Taking our results at face value, what might be the reasons for the different behaviour of household level inequality from the other series which seem to show declining inequality? It is possible to speculate that this divergence may well be due to a combination of factors, including: labour market participation decisions of the household and/or income “topping-up” by low wage households. It may be that these practices were more widespread and pervasive than is currently recognised in the existing scholarship on historic income inequality.

In more detailed work using the British survey data, Gazeley *et. al.* (2017) support the finding here of no evidence of declining inequality among working households in the first sixty years of the twentieth century. They find that this lack of movement was due to the offsetting effects of a variety of forces, some equalising and some tending to raise inequality. For some other European countries, it may also be possible to establish trends in inequality over time. But here, taking the evidence as a whole, it seems that for working households this was a broadly stable period in the European household income distribution.

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Table 1 Distribution of surveys used by country and time period

| | <i>Pre-1900</i> | <i>1900-1919</i> | <i>1920-1939</i> | <i>1940-1949</i> | <i>1950-1960s</i> |
|-------------|-----------------|------------------|------------------|------------------|-------------------|
| Austria | | 1 | 2 | | |
| Belgium | 4 | 2 | 1 | 1 | |
| Denmark | 1 | 1 | 3 | 2 | 1 |
| Finland | | 1 | 3 | | 3 |
| France | 1 | 4 | 1 | 5 | 4 |
| Germany | 2 | 8 | 5 | | 1 |
| Greece | | 1 | 1 | | 8 |
| Ireland | 2 | | | | 1 |
| Netherlands | | 3 | 3 | | 1 |
| Norway | | 6 | 1 | 1 | 7 |
| Poland | | | 3 | | 1 |
| Sweden | | 4 | 5 | 4 | 3 |
| Switzerland | 1 | 2 | 6 | | |
| UK | 1 | 1 | 1 | | 2 |
| Sum | 12 | 34 | 35 | 13 | 30 |

Table 2: Average 'naïve' Gini coefficients and decile ratios, by survey characteristics

| | N | Gini | 90/10 | 90/50 | 50/10 |
|-------------------------------------|----|-------|-------|-------|-------|
| <i>Derived from data:</i> | | | | | |
| Individual household level | 42 | 0.234 | 3.16 | 1.76 | 1.76 |
| Grouped by income/expenditure | 55 | 0.229 | 4.04 | 1.86 | 2.00 |
| Grouped by income and other aspects | 42 | 0.223 | 3.64 | 1.74 | 2.01 |
| Grouped by other aspects | 15 | 0.112 | 2.13 | 1.27 | 1.62 |
| <i>Era</i> | | | | | |
| Pre-1920 | 60 | 0.216 | 2.85 | 1.73 | 1.61 |
| 1920-1939 | 41 | 0.197 | 3.12 | 1.69 | 1.78 |
| Post-1939 | 54 | 0.235 | 4.54 | 1.80 | 2.31 |
| <i>Survey Design</i> | | | | | |
| Unknown | 64 | 0.202 | 2.84 | 1.68 | 1.64 |
| Non-random/non representative | 34 | 0.213 | 3.02 | 1.67 | 1.78 |
| Non-random/attempted representative | 19 | 0.190 | 3.18 | 1.62 | 1.92 |
| Stratified random sample | 26 | 0.276 | 5.51 | 2.01 | 2.56 |

Sources and notes: see text

Table 3: Average Beta-Lorenz Gini coefficients and decile ratios, by survey characteristics

| | N | Gini | 90/10 | 90/50 | 50/10 |
|-------------------------------------|----|-------|-------|-------|-------|
| <i>Derived from data:</i> | | | | | |
| Individual household level | 42 | 0.236 | 3.00 | 1.71 | 1.69 |
| Grouped by income/expenditure | 55 | 0.259 | 3.15 | 1.62 | 1.80 |
| Grouped by income and other aspects | 42 | 0.230 | 3.47 | 1.68 | 1.88 |
| Grouped by other aspects | 15 | 0.124 | 1.80 | 1.28 | 1.39 |
| <i>Era</i> | | | | | |
| Pre-1920 | 60 | 0.229 | 2.72 | 1.62 | 1.62 |
| 1920-1939 | 41 | 0.214 | 2.56 | 1.53 | 1.58 |
| Post-1939 | 54 | 0.249 | 3.85 | 1.71 | 2.03 |
| <i>Survey Design</i> | | | | | |
| Unknown | 64 | 0.215 | 2.48 | 1.55 | 1.57 |
| Non-random/non representative | 34 | 0.222 | 2.77 | 1.60 | 1.70 |
| Non-random/attempted representative | 19 | 0.199 | 2.52 | 1.49 | 1.64 |
| Stratified random sample | 26 | 0.295 | 4.28 | 1.85 | 2.17 |

Sources and notes: see text

Table 4. European Gini coefficient averages by sampling method (% , with number of coefficients in brackets)

| <i>Sampling method</i> | <i>Pre-1920</i> | <i>Interwar period</i> | <i>Post-1939</i> |
|------------------------|-----------------|------------------------|------------------|
| Unknown | 21.5 (35) | 18.5 (23) | 18.6 (6) |

Inequality among European Working Households, 1890-1960

| | | | |
|----------------------------|-----------|----------|-----------|
| Snowball sample | 20.6 (22) | 23.7 (8) | 20.4 (4) |
| Other non-random | | 23.0 (3) | 18.2 (16) |
| Stratified random sampling | | 21.6 (3) | 28.4 (23) |

Table 5. Understanding the variations in European working household inequality 1890-1960

| Naïve estimates of | Gini | 90/10 | 90/50 | 50/10 |
|--|-----------|----------|----------|---------|
| <i>Explanatory factors:</i> | | | | |
| <i>Era: (default=post 1939)</i> | | | | |
| Before 1920 | -0.011 | -0.376 | 0.005 | -0.174 |
| 1920-1939 | -0.005 | -0.252 | -0.005 | -0.093 |
| Post 1939*Stratified | 0.062 | 1.873 | 0.137 | 0.864* |
| Random | | | | |
| <i>Survey design</i> | | | | |
| Rural households included | 0.053*** | 1.24** | 0.339** | 0.310 |
| Urban households included | 0.068** | 2.19** | 0.268* | 0.864** |
| Blue-collar only | 0.002 | 0.318 | -0.039 | 0.263 |
| Not grouped by income/expenditure | -0.080*** | -0.60*** | -0.43*** | -0.080 |
| Grouped by income and other factors | -0.026 | -0.349 | -0.184* | -0.030 |
| Groups means estimated | 0.049*** | 1.19** | 0.126 | 0.50** |
| <i>Sampling method (default=unknown)</i> | | | | |
| Non-random, nonrepresentative | -0.017 | -0.36 | -0.007 | -0.261 |
| Non-random, representative | -0.024 | -0.19* | -0.133 | -0.903 |
| Stratified random sampling | -0.033 | -0.67 | 0.029 | -0.793 |
| R ² | 0.63 | 0.61 | 0.59 | 0.62 |
| N | 132 | 132 | 132 | 132 |

Notes: inequality statistics estimated from 132 surveys from 14 countries, using the simplest interpolation methods. Country controls included. *, **, *** signify conventional significance at the 10%, 5% and 1% levels.

Table 6. Understanding the variations in European working household inequality 1890-1960

| Beta-Lorenz Dep. Var.: | Gini | 90/10 | 90/50 | 50/10 |
|--|-------------|--------------|--------------|--------------|
| <i>Explanatory factors:</i> | | | | |
| <i>Era: (default=post 1939)</i> | | | | |
| Before 1920 | -0.013 | -0.045 | -0.017 | -0.005 |
| 1920-1939 | -0.007 | -0.055 | -0.031 | 0.008 |
| Post 1939*Stratified | 0.070 | 1.752** | 0.123 | 0.792*** |
| Random | | | | |
| <i>Survey design</i> | | | | |
| Rural households included | 0.055*** | 1.069*** | 0.213*** | 0.309** |
| Urban households included | 0.082** | 1.070** | 0.173 | 0.393* |
| Blue-collar only | -0.002 | 0.064 | -0.025 | 0.067 |
| Not grouped by income/expenditure | -0.094*** | -0.523 | -0.234 | -0.099 |
| Grouped by income and other factors | -0.044** | 0.088 | 0.022 | 0.003 |
| Groups means estimated | 0.060*** | 0.687** | 0.051 | 0.314*** |
| <i>Sampling method(default=unknown)</i> | | | | |
| Non-random, nonrepresentative | -0.015 | -0.305 | -0.068 | -0.083 |
| Non-random. representative | -0.022 | 0.395 | -0.180* | -0.010 |
| Stratified random sampling | -0.040 | -0.833 | -0.08 | -0.349 |
| R ² | 0.64 | 0.72 | 0.65 | 0.64 |
| N | 132 | 132 | 132 | 132 |

Notes: inequality statistics estimated from 132 surveys from 14 countries, using the simplest interpolation methods. Country controls included. *, **, *** signify conventional significance at the 10%, 5% and 1% levels.

Appendix 1: Data reference list

Austria

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Appendix 2: Testing alternative methods of estimating inequality from grouped data.

In order to sketch our methods here, we present some results from testing for the biases of different estimators of inequality measures. The testing approach employs bootstrap sampling. We repeatedly drew random subsamples from two different individual data sets, and from each subsample create a grouped data set this data is then tabulated with vary number of income groups created. Then for each set of grouped data, we make up to six different estimates of the Gini coefficient and various decile ratios using various estimation methods. We then compare these estimates with the inequality measure taken directly from the subsample of individual households, the average of the differences between these two measures across the 500 repetitions, is thus the bias of the estimators. The estimates were drawn for tabular data grouped by income ranging between 5 to 10 groups, as these are the most frequent numbers of groups in the tables we located.

Consider Table A1, which is taken from Gutierrez Rufrancos and Newell (2017). The data used to generate this table were from BMinistry of Labour (). The table provides the evaluation of six methods for income inequality estimation, as listed in the first column. Each cell can thus be thought of as our assessment of the bias associated with using this estimation technique. The ‘Groups Naïve Frequency Method’ simply assigns the group average income to each household in the group and then calculates the Gini coefficient using the usual unit record formulas⁶ Thus we find that this first method overestimates the Gini on this dataset by between one and two points on average.

The second row gives the bias, similarly defined, of using Hermite interpolation within the groups, , as suggested by Gastwirth and Glauber (1976). Gutierrez Rufrancos The third row reports the bias from a parametric approach, estimating lognormal parameters from an interval regression exploiting the upper and lower bounds of the groups. The fourth method simply estimates lognormal with an ordinal least squares regression on group means. The fifth method estimates Kakwani’s Beta Lorenz curve (Kakwani, 1980), and the sixth method combines Hermite-interpolation with Beta Lorenz estimation see Gutierrez Rufrancos and Newell *op. cit.* for a discussion.

⁶ In practice this was estimated using Jenkins (1999) module for unit record estimation in Stata - ineqdeco-.

Table A1: Bias on Estimates of Gini Coefficient using Income from Group data, UK 1953/4

| # of Bins (Groups) | 5 | 6 | 7 | 8 | 9 | 10 | Rank |
|---------------------------------------|-------|-------|-------|-------|-------|-------|------|
| Groups Naive Freq. Weighted | 0.010 | 0.014 | 0.017 | 0.018 | 0.019 | 0.020 | 1 |
| Hermite Interpolation (bands) | 0.041 | 0.038 | 0.036 | 0.034 | 0.033 | 0.031 | 5 |
| Lognormal interval regression (bands) | 0.014 | 0.017 | 0.019 | 0.021 | 0.023 | 0.024 | 2 |
| Lognormal OLS | 0.046 | 0.048 | 0.057 | 0.060 | 0.064 | 0.074 | 6 |
| Beta-Lorenz | 0.027 | 0.027 | 0.027 | 0.026 | 0.026 | 0.026 | 4 |
| Hermite-Beta Lorenz (bands) | 0.044 | 0.033 | 0.025 | 0.021 | 0.018 | 0.014 | 3 |

Notes to Table 1. Each cell gives the mean bias from 100 bootstrap replication of each group data estimation method, by the assumed number of bins or groups.

The last column of Table A1 ranks the estimators inversely with the size of the mean bias across the different number of group sizes. Across both datasets, Gutierrez Rufrancos and Newell op. cit. find the least-biased estimator to be the Beta-Lorenz. It characterises the decile ratios very well. It does not, however, provide the best estimate of the Gini coefficient. Where the data only provides interval information, the best estimator is the combination of the Beta-Lorenz and the Hermite interpolation. However, in some extreme cases this fails to resolve numerically the non-linear least squares. When this is the case, the suggested second-best performer is the interval regression based lognormal estimator. As a result, eleven of our twenty-five inequality estimates based on grouped data are estimated by Hermite-interpolated-Beta Lorenz, eight are estimated by Beta-Lorenz and six by the interval based lognormal estimator.